. I : I D S ^{instytut} badań strukturalnych

IBS WORKING PAPER 04/2019 REVISED VERSION JANUARY 2021

TECHNOLOGY, SKILLS, AND GLOBALISATION: EXPLAINING INTERNATIONAL DIFFERENCES IN ROUTINE AND NONROUTINE WORK USING SURVEY DATA

Piotr Lewandowski Albert Park Wojciech Hardy Yang Du Saier Wu

TECHNOLOGY, SKILLS, AND GLOBALIZATION: EXPLAINING INTERNATIONAL DIFFERENCES IN ROUTINE AND NONROUTINE WORK USING SURVEY DATA*

Piotr Lewandowski*, Albert Park*, Wojciech Hardy*, Yang Du*, Saier Wu°

Abstract

The shift away from manual and routine cognitive work, and towards non-routine cognitive work is a key feature of labor markets. There is no evidence, however, if the relative importance of various tasks differs between workers performing seemingly similar jobs in different countries. We develop worker-level, survey-based measures of task content of jobs - non-routine cognitive analytical and personal, routine cognitive and manual - that are consistent with widely-used occupation-specific measures based on O*NET database. We apply them to representative surveys conducted in 47 countries at different stages of development. We find substantial cross-country differences in the content of work within occupations. Routine task intensity (RTI) of jobs decreases significantly with GDP per capita for high-skill occupations but not for middle- and low-skill occupations. We estimate the determinants of workers' RTI as a function of technology (computer use), supply of skills, globalization (specialization in global value chains) and sector structures, and decompose their role in accounting for the variation in RTI across countries. Computer use, better education, and higher literacy skills are related to lower RTI. Globalization (as measured by sector foreign value-added share) increases RTI in poorer countries, especially in low-skilled occupations. Differences in technology endowments and in skills' supply matter most for cross-country differences in RTI, with globalization also important. Technology and skills contribute the most to the differences in RTI among workers in high-skilled occupations; while globalization contributes the most to differences among workers in low-skilled occupations.

Keywords: task content of jobs, deroutinisation, global division of labor, PIAAC, STEP, CULS.

JEL: J21, J23, J24

[•] We thank Peng Jia, Cheng Jie and Zeyang Yu for their excellent research assistance, Emily Pawlowski of the American Institutes for Research for help with using the US PIAAC restricted data. We thank participants of ASSA2019 in Atlanta, EALE2018 in Lyon, Jobs and Development Conference 2018 in Bogotá, TASKS V in Bonn, seminars at UCSD, The World Bank, and Princeton University, and workshops in Beijing, Geneva, Luxembourg and Warsaw for their insightful comments. We acknowledge support from the Hong Kong Research Grants Council General Research Fund [16504217]. The usual disclaimers apply. All errors are our own.

[•] Institute for Structural Research (IBS), Warsaw, and IZA, Bonn. E-mail: piotr.lewandowski@ibs.org.pl. Corresponding author.

[♥] HKUST Institute for Emerging Market Studies, Hong Kong, and IZA, Bonn. E-mail: albertpark@ust.hk.

A Institute for Structural Research (IBS), Warsaw, and Faculty of Economics, University of Warsaw. E-mail: wojciech.hardy@ibs.org.pl.

Institute of Population and Labor Economics, Chinese Academy of Social Science (CASS), Beijing, and IZA, Bonn. E-mail: duyang@cass.org.cn.
[°] HKUST, Hong Kong. E-mail: saier.wu@connect.ust.hk.

1. Introduction

In recent years, there has been an explosion of interest in how the future of work will be shaped by new technologies. Much speculation focuses on whether and how quickly specific job tasks can be performed by robots or artificial intelligence (Arntz et al., 2017; Frey and Osborne, 2017; McKinsey, 2017). Economists have found analysis of changes in job tasks to be a fruitful way to understand how labor market outcomes are influenced by the changing nature of work (Acemoglu and Autor, 2011; Autor, 2013; Firpo et al., 2011). In the US and other advanced countries, the relative share of routine jobs – both cognitive and manual – has declined over time, presumably because such jobs are easily replaced by computers or automation, or can be outsourced to other countries (Autor et al., 2003; Goos et al., 2014; Jensen and Kletzer, 2010; Michaels et al., 2014; Spitz-Oener, 2006). Because these jobs tend to be middle-skill jobs, the decline of routine jobs has contributed to wage polarization by hollowing out the middle of the wage distribution. However, in developing countries and emerging markets, evidence on how the nature of work is changing is decidedly mixed (World Bank, 2019). For example, there is evidence that in China and some Central Eastern European countries, the role of routine-intensive occupations actually increased in recent decades (Du and Park, 2018; Hardy et al., 2018).

In this paper, we study how four fundamental forces predict the task content of jobs across countries: technology, globalization, structural change, and supply of skills.¹ Large labor productivity differences across countries as well as significant differences in information and communication technologies (ICT) and robot adoption suggest that large technological gaps remain across countries (Eden and Gaggl, 2020; Hsieh and Klenow, 2010). Globalization is expected to lead to the outsourcing of routine-intensive tasks from high-wage countries to low-wage countries (Grossman and Rossi-Hansberg, 2008; Hummels et al., 2018). Structural changes such as industrialization and the growth of services alter the demand for goods and services which alter the demand for different types of jobs (Bárány and Siegel, 2018). Finally, the labor force in poorer countries often is much less educated, which could influence the optimal assignment of routine and non-routine tasks (World Bank, 2019).

The analysis of task demand in the US and globally has been greatly facilitated by the codification of the task content of different occupations in the US by the Department of Labor, first through the Dictionary of Occupational Titles (DOT) dating back to 1939, and since 2003 through the Occupation Information Network (O*NET). These databases provide detailed and periodically updated descriptions of the specific tasks associated with each occupation in the US. Acemoglu and Autor (2011) used O*NET data to construct what have now become standard indices of different types of job tasks: non-routine cognitive analytical, non-routine cognitive interpersonal, routine cognitive, routine manual, and non-routine manual. Because other countries have not systematically collected similar information on occupational job tasks, analyses of task demand in other countries has frequently used the US O*NET task data, requiring the assumption that the task content of occupations in those countries is identical to the US (Arias et al., 2014; Goos et al., 2014; Hardy et al., 2018; Lewandowski et al., 2020). This is almost certain to be problematic for less developed countries, given significant heterogeneity in the four fundamental forces described above.

¹ We focus on factors that directly influence prices of outputs and factors, and firms' technology. We do not consider institutional factors, although we recognize their importance in shaping technology, globalization, structural change, and skills, as well as the organization of work (firm size, management structure, etc.) all of which may influence job tasks.

In this paper, we present new evidence on the global distribution of tasks, focusing on the determinants of crosscountry differences in the nature of work. We use micro-data on job tasks collected from large-scale surveys of workers in 47 countries around the world spanning developed and developing countries. The survey data come from three sources: the OECD's Programme for the International Assessment of Adult Competencies (PIAAC), the World Bank's Skills toward Employment and Productivity (STEP) surveys, conducted in middle- and low-income countries, and the China Urban Labor Survey (CULS) which included a module based on STEP. We develop harmonized survey-based measurements of non-routine cognitive analytical, non-routine cognitive interpersonal, routine cognitive, and manual tasks which closely mirror widely used task measurements for occupations proposed by Acemoglu and Autor (2011) based on the US O*NET data. However, our measures are worker-specific, enabling us to capture within-occupation differences in job tasks both within and across countries. Even in the US, research has shown considerable variation in tasks among workers within the same occupation (Autor and Handel, 2013). Construction of worker-specific task measures that are consistent with O*NET and cover low-, middle-, and highincome countries is a first main contribution of this study.²

Our second contribution is to document new stylized facts about cross-country differences in the task content of jobs. On average, workers in the more developed economies perform more non-routine cognitive tasks, both analytical and interpersonal, and less routine tasks. The relationship between routine task intensity (a summary measure that combines routine cognitive and non-routine cognitive tasks) and country GDP per capita differs quite markedly for different occupation groups. In high-skilled occupations (e.g., managers, professionals), there is a sharp gradient with respect to GDP per capita, with work being more routine-intensive in poorer countries. However, for middle-skill occupations like clerical workers, and low-skill occupations like plant and machine operators and assemblers, we find an inverse-U shaped relationship between the relative routine-intensity of tasks and country's level of development. Overall, cross-country differences in task content within the same occupations are sizable. Understanding the extent and nature of these differences is of both scientific and policy relevance, as they reflect differences in the nature of work that can inform future labor market challenges, such as the share of jobs that can be automated (Arntz et al., 2017; Frey and Osborne, 2017).

Our third and most important contribution is to quantify for the first time how four fundamental forces – technology, globalization, supply of skills, and structural change – are associated with cross-country differences in the task content of jobs. Previous research has documented associations between specific factors for subsets of countries, but this study is the first to examine the impact of all of these factors in a comprehensive framework, and for countries that span low-, middle-, and high-income countries, and using survey-based measures.³ All of the previous studies assume that tasks within occupations are identical across countries.

² Other studies of tasks that use international survey data typically focus on less diversified samples of countries: (de la Rica et al., 2020) and (Marcolin et al., 2019) used PIAAC data to study OECD countries, while (Lo Bello et al., 2019) used STEP data to cover low-income countries).

³ For example, earlier research documents the importance of ICT technology for the demand for tasks in the OECD countries (Akerman et al., 2015; Autor et al., 2006, 2003; Deming, 2017; Spitz-Oener, 2006), but no studies document the relationship between ICT and tasks in a cross-country setting that includes low-income countries. Evidence also exists that offshoring contributes to the shift away from routine work in the OECD countries (Goos et al., 2014; Hummels et al., 2018; Oldenski, 2012) and that participation in global value chains leads to a higher share of routine-intensive occupations in some developing

Our assessment of the relative importance of the four fundamental factors in predicting cross-country task differences starts with a regression in which workers' routine task intensity is regressed on individual, sector, and country-level variables that capture the four factors, for the pooled, global sample of all workers. Technology is captured by country-sector computer use calculated from the survey data, as well as country / sector robot and ICT stocks, globalization is measured by the foreign share of value added in the country-sector plus this share interacted with GDP per capita, structural change is captured by 18 sector indicator variables and GDP per capita, and skills are captured by individual education, demographics (age, sex), and a direct test of literacy proficiency which allows us to measure skill more accurately than studies which only use data on education attainment (which cannot capture education quality differences). Given our finding that GDP per capita predicts relative routineness of tasks differently for different occupation groups, we run separate task regressions for workers in high-skilled occupations (managers, professionals and technicians), middle-skilled occupations (clerks, sales and services workers) and low-skilled occupations (craft and related trades workers, plant and machine operators and assemblers, elementary occupations).

Using the coefficients from these regressions, we decompose gaps in mean routine task intensity across countries into the contributions associated with the four fundamental forces. We decompose the cross-country variance in mean routine task intensity, and compare each country to the US, summarizing the importance of different factors in predicting gaps between the US and groups of countries sorted by GDP per capita.

Our regression-based decompositions show that technology, the supply of skills and globalization are all strongly associated with cross-country differences in routine task intensity (RTI). International differences in technology use are especially important in accounting for cross-country variation in RTI in high-skilled occupations that are typically rich in non-routine tasks, highlighting the complementarity between non-routine tasks and ICT (Autor et al., 2003). Globalization contributes the most to cross-country differences in RTI among workers in low-skilled occupations. This finding is in line with the view that offshoring enables countries to specialize, within industries, according to their abundant factors (Grossman and Rossi-Hansberg, 2008). We also find that the supply of skills contributes to the cross-country differences in tasks mainly by shaping the employment shares of high-, middle-, and low-skilled occupations. Moreover, in the low- and middle-income countries, lower supply of skills accounts for a large share of the difference in RTI compared to the US. Although the complementarity of supply of skills and technology use has been acknowledged as a key factor behind cross-country differences in technology use (Eden and Gaggl, 2020) it is often overlooked in the studies of tasks that are focused on the most developed countries. We provide evidence that it in the poorer countries it should be accounted for as it may help to understand not only why the shares of high-skilled occupations are low, but also why the tasks performed by workers in given occupations are more routine-intensive. Finally, we show that differences in occupational structure across countries account for a relatively small share of cross-country differences in task content. This highlights the importance of using comparable survey data to accurately estimate the extent of cross-country differences in task content and the determinants of those differences.

countries (Reijnders and de Vries, 2018). Regarding skill supply, a positive relationship between the supply of tertiary educated workers and non-routine tasks has been documented by studies using O*NET data (Hardy et al., 2018; Montresor, 2019; Salvatori, 2018). Structural change has been identified as relevant for polarization and shifts in tasks over time, both theoretically (Bárány and Siegel, 2018) and empirically (Du et al., 2017; Hardy et al., 2018).

In the second section we outline our methodology of creating the task content measures using the US PIAAC and O*NET data, and applying them to 47 countries covered by the PIAAC, STEP and CULS surveys. In the third section we present the cross-country differences in task structures, and in the fourth section we examine the determinants of these differences. The fifth section concludes.

2. Data and Task Measurement

2.1 Data for measurement of job tasks

Our aim is to create task content measures based on PIAAC and STEP surveys which are worker-specific but are as consistent as possible with well-established measures of job tasks. To accomplish this objective, we first use the US PIAAC dataset to create measures that maximize consistency with US O*NET occupation-specific task measures popularized by Acemoglu and Autor (2011).

We use data from three comparable surveys: OECD's Programme for the International Assessment of Adult Competencies (PIAAC) the World Bank's Skills Measurement Program (STEP), and the third wave of the China Urban Labor Survey (CULS) conducted by the Institute of Population and Labor Economics of the Chinese Academy of Social Science (CASS). Our sample covers 47 countries in total.

In two rounds of PIAAC surveys (in 2011-12 and 2014-15), data was collected in 37 countries that made their data publicly available.⁴ The countries covered by PIAAC are high- or middle-income countries (see Appendix A for the list of countries). The survey respondents were aged 16-65, with sample sizes ranging from about 4000 in Russia to 26000 in Canada.⁵ STEP surveys are available for 13 low- or middle-income countries, out of which we use nine (Appendix A).⁶ The surveys were conducted between 2012 and 2014 of urban residents aged 15-64, with sample sizes ranging from about 2400 (in Ukraine) to approx. 4 000 (in Kenya), of urban residents aged 15-64.⁷ We also use the third wave of CULS which included the "skill use at work" questionnaire of STEP and therefore it is directly comparable to STEP. The survey was conducted in 2016 in six large cities in China (Guangzhou, Shanghai, and Fuzhou on the coast, Shenyang in the northeast, Xian in the northwest, and Wuhan in central China) and has a sample of 15 448 individuals.⁸ We refer to CULS as one of the STEP countries.⁹

⁴ In the US, PIAAC was supplemented by an additional wave aimed at enhancing the sample size, while retaining representativeness. We use this sample which is available from the US National Center for Education Statistics (NCES).

⁵ Individuals aged 15-year were also surveyed in Australia and Chile. Individuals aged 66-74 were surveyed in Australia.

⁶ We decided against using three available STEP datasets: Yunnan (China), Sri Lanka, Ukraine, and Vietnam. For China, we use the CULS data instead of the STEP survey for the Chinese Yunnan province, as the former contains far more observations and covers a more comprehensive area. Yunnan is one of the poorer and more rural provinces in China so it might not reflect the dominant patterns of work in Chinese urban areas. The survey of Sri Lanka includes too few observations in urban areas (about 650 workers), the Ukraine survey lacks one of the questions required for our task measures, the Vietnam survey has low quality of data on skill use at work.

⁷ Because nearly all STEP surveys were urban only, for Laos which surveyed both urban and rural residents we drop the rural part of sample in order to ensure consistency.

⁸ The survey sampled 260 neighborhoods, 2 581 migrant households and 3 897 local households.

⁹ We reweight the STEP and Indonesian⁹ data in order to achieve representativeness of the occupational structures in urban areas. To this aim, we retain the original shares of workers in agriculture and elementary occupations and adjust the

2.2 Task measures' definitions based on the survey data

To construct survey-based task measures consistent with those based on O*NET, we first identify harmonized survey questions available in both PIAAC and STEP surveys whose content is similar to the questions used by Acemoglu and Autor (2011) to construct the O*NET-based task measures. Then we systematically search for combinations of appropriate survey questions (and best groupings of answers) for which US PIAAC survey-based measures (averaged for each occupation) are most correlated with US O*NET-based occupation measures. Because PIAAC and STEP include only one question on physical tasks, we apply our procedures to the cognitive tasks measures only. For methodological details, see Appendix B.

Our procedure results in the following survey-based task definitions. The non-routine cognitive analytical task measure is based on questions on solving problems, reading news, reading professional journals, solving problems and programming. The non-routine cognitive interpersonal task measure is based on supervising others and making presentations. The routine cognitive task measure is based on the ability to change the order of tasks (reversed, so not being able to change the order of tasks), filling out forms, and making speeches or giving presentations (reversed, so making no speeches and giving no presentations). The manual task measure is based on the item describing if a job usually involves working physically for a long period. The cutoffs for each item are presented in Table 1.

In the US, our survey-based measures follow closely the task measures based on O*NET. At the 3-digit occupation level, the correlations between the survey measures (occupation-level averages) and the Acemoglu and Autor (2011) measures range from 55% (routine cognitive) to 77% (non-routine cognitive analytical, Table 1).¹⁰

We use the definitions presented in Table 1 to calculate worker job task content measures for all countries studied.¹¹ We also merge O*NET with PIAAC, STEP and CULS in order to calculate the Acemoglu and Autor (2011) task measures for each country. For both measures, we standardize the measure using the relevant mean and standard deviation in the US. Hence, for each task measure, zero reflects the US average and 1 reflects standard deviations in the US. As the STEP surveys are urban surveys, we omit skilled agricultural workers (ISCO 6) in all countries to improve comparability.

distribution of other 1-digit ISCO occupations in line with occupational distributions reported in the International Labour Organization Database (ILOSTAT). In the case of China, we use the urban occupational distribution from the 2015 Census to reweight the CULS data.

¹⁰ The highest correlations obtained at the 4-digit occupation level range from 62% to 79%.

¹¹ The Ukrainian STEP does not include the question about reading of professional items so we are unable to calculate the non-routine cognitive analytical measure for Ukraine. We use only the non-routine cognitive personal measure in the RTI.

Task content	Non-routine cognitive analytical	Non-routine cognitive interpersonal	Routine cognitive	Manual
Task items	Solving problems Reading news (at least once a month – answers 3,4,5) Reading professional journals (at least once a month – answers 3,4,5) Programming (any frequency – answers 2,3,4,5)	Supervising others Making speeches or giving presentations (any frequency - answers 2,3,4,5)	Changing order of tasks - reversed (not able) Filling out forms (at least once a month – answers 3,4,5) Making speeches or giving presentations - reversed (never)	Physical tasks
Correlation with Acemoglu and Autor (2011) measures	0.77	0.72	0.55	0.74

Table 1. The task items selected to calculate task content measures with the US PIAAC data

Note: The cuttoffs for the "yes" dummy in brackets. The full wording of questions and definitions of cutoff are presented in Appendix C. Source: own elaboration based on US PIAAC and O*NET data.

We create a synthetic measure of relative routine task intensity (RTI) at a worker-level, using the formula:

$$RTI = \ln(r_{cog}) - \ln\left(\frac{nr_{analytical} + nr_{personal}}{2}\right), \tag{1}$$

)

whereby r_{cog} , $nr_{analytical}$ and $nr_{personal}$ are routine cognitive, non-routine cognitive analytical and non-routine cognitive personal task levels, respectively. For each task, we add the lowest score in the sample to the scores of all individuals, plus 0.1, to avoid non-positive values in the logarithm. This definition follows the literature (Autor and Dorn, 2009, 2013; Goos et al., 2014) but we omit the manual tasks for two reasons: first, we cannot distinguish between routine and non-routine manual tasks. Second, the manual measure is less comparable across countries than the other measures – we provide evidence in the next section. For consistency, we standardize the RTI using its mean and standard deviation in the US.

2.3 Econometric methodology

To shed light on factors associated with cross-country differences in routine task intensity, we estimate pooled OLS regressions of the form:

$$RTI_{ijsc} = \beta_0 + \beta_1 Z_{sc} + \beta_2 G_{sc} + \beta_3 E_{ijsc} + \lambda_s + \varepsilon_{ijsc}, \qquad (2)$$

Here, RTI_{ijsc} is the routine task intensity of individual *i* in occupation *j* in sector *s* in country *c*, Z_{sc} is technology used in sector *s* in country *c*, G_{sc} measures globalization in sector *s* in country *c*, E_{ijsc} are the individual skills of workers, and λ_s are sector fixed effects.

We also expand the benchmark specification by adding occupation fixed effects, τ_o :

$$RTI_{ijsc} = \beta_0 + \beta_1 Z_{sc} + \beta_2 G_{sc} + \lambda_s + \beta_3 E_{ijsc} + \tau_o + \varepsilon_{ijsc}, \tag{3}$$

Because the regressions are cross-sectional, they are best thought of as characterizing equilibrium allocations of tasks rather than being interpreted causally. The technology, globalization, and structural change measures are all country-sector level measures, which are plausibly exogenous to the decisions of individual firms and workers. We

measure skills at the individual level, given that education and literacy are mostly pre-determined before entering the labor market. We have also conducted the analysis defining skill levels at the sector-level and discuss how this alters the results below.

Turning to measurements, the main technology variable is the share of workers in sector *s* in country *c* who use computers at work. PIAAC and STEP surveys include a question on individual computer use, and we aggregate this variable to the sector level due to concerns that decisions about computer use and tasks are made simultaneously. We use a quadratic specification to allow for a potential non-linear relationship between computer use and RTI. Separately, we also test the impact of robot stock per worker by sector (International Federation of Robotics data), and country-level ICT capital stock per worker (Eden and Gaggl, 2020). Adding these variables turns out not to alter the main findings in an important way, but these data are available for only 32 countries so we exclude them from our preferred specification.

We employ two variables to measure globalization -- participation in global value chains (henceforth GVC participation, Wang et al., 2017), and foreign direct investment (FDI) stock as a share of GDP.¹² Our basic GVC participation variable is the backward linkage-based measure defined as the foreign value added share in production of final goods and services (FVA share). For robustness we also use the forward linkage-based measure (domestic value added from production of intermediate exports or domestic factor content in intermediate exports, Wang et al., 2017). We allow for different effects of GVC participation in developed and developing countries by interacting the GVC participation with GDP per capita (log, demeaned). This captures the prediction that globalization reduces routine tasks in rich countries and increases them in poor countries.

To measure worker skills, we include a test-based measure of literacy skills (four proficiency levels), education level (primary, secondary, tertiary), age (measured by 10-year age groups), and gender. The literacy test is comprehensive and quantifies individuals' skills to understand, evaluate, use and engage with written texts in personal, work-related, societal and educational contexts (PIAAC Literacy Expert Group, 2009).¹³

To capture the impact of structural change, we add indicator variables for 18 of 19 sectors based on the one-digit codes of International Standard Industrial Classification (ISIC Rev.4), as well as their interactions with GDP per capita (log, demeaned).¹⁴

¹² Data sources and precise definitions of the technology and globalization variables are provided in Appendix E.

¹³ We account for the fact that PIAAC and STEP include multiple plausible values of the literacy proficiency variables. To this aim, we use the "pv" package in Stata that implements the (Rubin, 1987) combination methods which are commonly used in the skill assessment literature. China and Laos did not collect literacy data, so we impute the literacy scores for those two countries using a regression estimated for other countries, controlling for education, demographic characteristics, occupation and sector of employment, computer use at work as well as macroeconomic variables (GDP per capita, export, FDI).

¹⁴ In order to achieve consistent sector definition across all countries, we merge sectors D (Electricity, gas, steam and air conditioning supply) with E (Water supply; sewerage, waste management and remediation activities), and M (Professional, scientific and technical activities) with N (Administrative and support service activities).

We estimate the regressions for all workers, and for subgroups of workers.¹⁵ Given the evidence presented above that the cross-country patterns in routine-intensity vary for occupations of different skill levels, we estimate separate regressions for workers in high- (ISCO 1-3), middle- (ISCO 4-5) and low-skilled (ISCO 7-9) occupations.

In order to assess the relative importance of the four fundamental factors in predicting cross-country differences in tasks, we use the estimated coefficients to calculate a linear prediction of routine task intensity at the country level \overline{RTI}_c . We decompose the variance of RTI using the covariance-based decomposition proposed by (Morduch and Sicular, 2002). Formally, the contribution of a variable group, k, to the variance of RTI is defined as follows:

$$\sigma_{k} = \frac{cov(\beta_{k}\bar{x}_{c}^{k},\overline{RTI}_{c})}{var(\overline{RTI}_{c})},$$
(4)

We use the average worker characteristics in each country (denoted by dashed symbols) to decompose the difference in the linear prediction of RTI in country c, \widehat{RTI}_c , and the US, \widehat{RTI}_{US} , to the contributions of various factors:

$$\widehat{RTI}_{c} - \widehat{RTI}_{US} = \beta_{1} \left(\overline{Z_{\iota JSC}} - \overline{Z_{\iota JSUS}} \right) + \beta_{2} \left(\overline{G_{sc}} - \overline{G_{sU}} \right) + \lambda \left(\overline{S_{sc}} - \overline{S_{sUS}} \right) + \beta_{3} \left(\overline{E_{\iota JSC}} - \overline{E_{\iota JSUS}} \right), \quad (5).$$

For presentation purposes, we aggregate countries to three groups based on their development level, and use unweighted averages of differences in RTI, all explanatory variables and all contributions (Table 2).

Low- and Middle-Income Countries	Bottom High-Income Countries	Top High-Income Countries	Reference country
Kenya, Ghana, Lao, PDR, Bolivia, Indonesia, Ecuador,	Chile, Hungary, Poland, Lithuania, Slovakia, Cyprus,	France, Israel, Japan, New Zealand,	United States
Peru, China, Armenia,	Estonia, Greece, Czech Rep.,	United Kingdom, Belgium,	
Georgia, Colombia, Mexico, Kazakhstan, Russia, Serbia,	Slovenia, Spain, Korea, Rep., Italy	Germany, Canada, Finland, Austria, Netherlands,	
Turkey		Ireland, Sweden, Denmark, Norway, Singapore	

Source: own elaboration based on World Bank data.

3. Comparison of survey and O*NET task measures

The occupational patterns of the survey task content measures are mostly consistent with those based on the O*NET database. Both measures show that workers in high-skill occupations (ISCO 1-3) perform, on average, less routine-intensive tasks, while workers in middle- and low-skill occupations (ISCO 4-5, ISCO 7-9) perform more routine-intensive tasks (Figure 1). Note that although the survey-based RTI measure does not include manual tasks, it successfully captures the general routine aspect of work – not only its cognitive dimension. In particular, the routine task intensity among plant and machine operators and assemblers (ISCO 8), who perform highly routine jobs according to the RTI based on O*NET that accounts for manual routine tasks, is also high according to the survey measure. The one distinction is that the survey measure also shows that tasks performed by sales and

¹⁵ We drop Macedonia from our regression sample due to lack of data on globalization variables, and estimate our models on a sample of 46 countries.

services workers (ISCO 5) around the world are on average slightly more routine than tasks than the tasks performed by clerical support workers (ISCO 4, a reference group), contrary to O*NET measures.

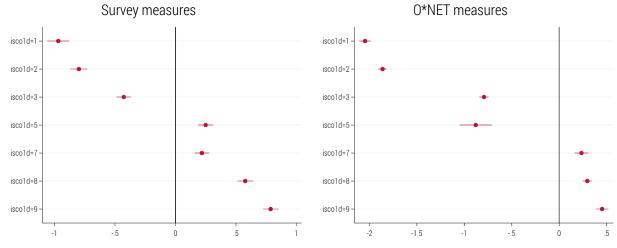


Figure 1. The differences in RTI across 1-digit ISCO occupations according to survey- and O*NET measures.

Note: coefficients pertaining to occupation fixed effects (1-digit ISCO) estimated in a worker-level model on RTI against occupation fixed effects and country fixed effects. Manual tasks are included in the RTI based on O*NET. Sample size 168,639. Reference groups: Clerical support workers (ISCO 4), the United States. Source: own calculations based on PIAAC, STEP, CULS and O*NET data.

The survey measures also exhibit pair-wise correlations that are consistent with those exhibited by Acemoglu and Autor (2011) measures (Table 2). The non-routine cognitive measures are strongly and positively correlated with each other, and negatively correlated with the routine cognitive and manual measures.¹⁶ The moderate positive correlation between the routine cognitive and manual measures is also very close to those calculated using the O*NET-based measures. This supports the validity of the survey-based RTI measure despite it lacking the manual tasks, as the other three task measures are enough to depict the occupational task sets.

	Non-routine cognitive analytical	Non-routine cognitive personal	Routine cognitive	Manual
	Acemoglu and Autor (201	1) measures based on O*	NET	
Non-routine cognitive analytical	1			
Non-routine cognitive personal	0.71	1		
Routine cognitive	-0.35	-0.54	1	
Manual	-0.64	-0.55	0.32	1
	Survey measures	based on US PIAAC		
Non-routine cognitive analytical	1			
Non-routine cognitive personal	0.64	1		
Routine cognitive	-0.49	-0.57	1	
Manual	-0.57	-0.58	0.42	1

Table 3. Pair-wise correlations between particular task content measures across 3-digit ISCO occupations in the US

Note: correlations between occupation-level averages in the case of survey measures. Weighted by employment level at the 3-digit ISCO level. Source: own calculations based on PIAAC and O*NET data.

¹⁶ This should alleviate the concerns related to the use of "Making speeches or giving presentations" variable in both the nonroutine cognitive personal measure and the routine cognitive measure. The negative correlation (across occupations) between them is virtually identical to the one implicitly implied by the Acemoglu and Autor (2011)measures.

The key difference between the survey measures and the measures based on O*NET is that the cross-country differences in task content within broad occupation groups are noticeably higher using the survey-based measures, while the differences between occupation-specific averages within particular countries are smaller.

The occupation-level averages of survey measures vary less between occupations than the measures based on O*NET (Figure B1 in Appendix B). At the 3-digit occupation level in the US, the standard deviations of tasks range from 0.50 (routine cognitive) to 0.67 (non-routine cognitive analytical) while the standard deviations of the O*NETbased tasks range from 1.02 (non-routine cognitive personal) to 1.23 (routine cognitive).¹⁷ The reason is that the O*NET measures are defined at the occupation level (4-digit ISCO), while the survey measures allow heterogeneity within occupations. At the 3-digit ISCO level, the within-group variance contributes from 65% (non-routine cognitive analytical) to 70% (non-routine cognitive personal, routine cognitive) to 83% (manual) of the overall variance of the survey measures in the US. In the sample of all countries, the RTI differences between occupations are also lower according to the survey measures than according to the O*NET measures (Figure 2).

Importantly, the cross-country differences in occupational RTI are generally higher according to the survey-based measure than according to the 0*NET-based measure. The standard deviations equal 0.23 vs. 0.11 among workers in high-skill occupations, 0.24 vs. 0.59 among workers in middle-skill occupations, and 0.21 vs. 0.11 among workers in low-skill occupations. Using the survey-based measures, differences may result from country differences in work even within the most narrowly defined occupations. Using the 0*NET-based measures, cross-country differences are driven entirely by cross-country differences in occupational structures at finer ISCO levels. To highlight this point, we use the US PIAAC survey data to construct mean task content measures at the same occupational level that we use for 0*NET measures. We then apply these to all workers assuming that job tasks in a specific occupations are identical across the world (Figure 2). The resulting task measures based on the assumption that occupations are identical across the world (but using our survey-based measure) follow the measures based on 0*NET much more closely than they follow country-specific measures based on the survey data. Hence, it is the country-specific measurement that matters.¹⁸

¹⁷ The high standard deviation of routine cognitive tasks based on O*NET is driven by negative outliers: occupations 521 (Street and Market Salespersons), 951 (Street and Related Services Workers) and 952 (Street Vendors, excluding food). If these outliers are ignored, the standard deviation of routine cognitive tasks turns out the lowest among the O*NET based measures (0.97), similarly to our measures.

¹⁸ The cross-country differences in O*NET-based measures may also result from inconsistencies in coding of occupations. Indeed, the O*NET measures suggest that in many low- or middle-income countries (e.g. Bolivia, Ghana, Indonesia, Mexico) workers in middle-skilled occupations perform highly non-routine work that on average is less routine-intensive than the work in high-skilled occupations (Figure 2). This implausible conclusion results from the fact that large shares of workers in these countries are classified as street sellers or services workers. These occupations require a lot of interpersonal tasks in the US but may not require as many of them in poorer countries. Indeed, the survey measures show that work in the middle-skilled occupations in poorer countries is quite intensive in routine tasks.

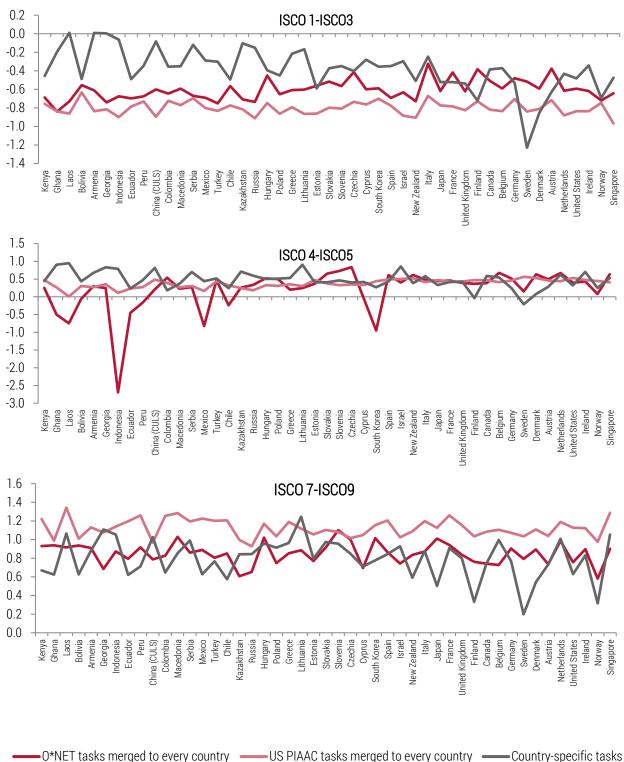
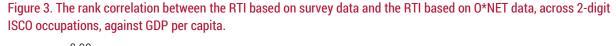
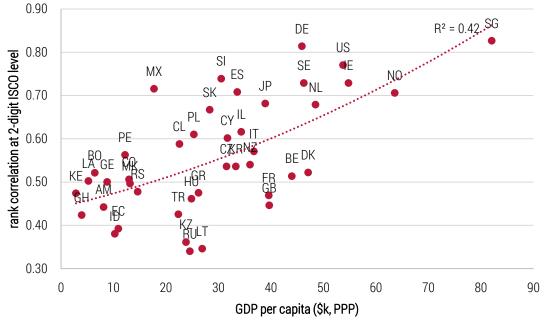


Figure 2. The average values of relative routine intensity (RTI) according to different methodologies, by occupational categories.

Note: Countries are ranked according to the GDP per capita level. For brevity, we aggregate occupations into three categories: highskilled (ISCO 1-3), middle-skilled (ISCO 4-5) and low-skilled (ISCO 7-9). Results for particular occupations are available upon request. Source: own calculations based on PIAAC, STEP, CULS, O*NET (tasks), and World Bank data (GDP). Finally, the higher is the development level of a country, the more similar are the occupational rankings based on the country-specific RTI and on the O*NET RTI. The within-country rank correlations of the survey-based and O*NET RTI measures (at 2-digit ISCO occupation level) are positively correlated with the GDP per capita levels (Figure 2). In the poorest countries in our sample the rank correlations amount to 0.4-0.5, while in the richest countries they amount to 0.7-0.8. Overall, our evidence shows that the task content of occupations differs with the level of development, and as a result the occupational characteristics defined in richer countries are not directly transferrable to poorer countries.





Source: own calculations based on PIAAC, STEP, CULS and O*NET data.

Descriptive evidence on the cross-country differences in the task content of jobs

We find substantial cross-country differences in the average values of particular task content measures. In general, the more developed countries exhibit higher average values of non-routine tasks than the less developed countries (Figure 4). The Nordic countries (Denmark, Sweden, Norway, Finland), most of the English-speaking countries (Canada, New Zealand, the UK and the US) and Singapore stand out with the highest levels of non-routine cognitive tasks. Perhaps not surprisingly, the less developed countries – Ecuador, Georgia, Ghana, Kazakhstan, Laos, Mexico, Colombia, Peru, Turkey, Indonesia, but also Lithuania and Greece – have the lowest average values of non-routine cognitive tasks. The average value of non-routine cognitive tasks, especially of analytical tasks, is also low in China. The differences between the average values of non-routine tasks in the highest-scoring and the lowest-scoring countries are of a magnitude comparable to a one standard deviation of particular task content values among the US workers.

The relationship between routine cognitive tasks and the level of development is inverse U-shaped (Figure 4). The least developed countries, the Nordic countries, Japan, and Austria exhibit the lowest values of routine cognitive tasks. On the other hand, Central, Eastern and Southeastern European countries (Lithuania, Czechia, Russia, Serbia, Slovakia, Slovenia) have the highest average values of routine cognitive tasks. The values of routine cognitive tasks are also high in Southern European countries (Greece, Italy), in the United Kingdom and Ireland.

The average values of manual tasks decrease somewhat with the level of development, but the relationship is less pronounced than in the case of non-routine cognitive tasks (Figure 4). However, we are able to use only one survey question to measure the manual task content, and it appears that this measure is less reliable than the other measures. Only five middle-income countries exhibit a higher value of manual tasks than the United States, and some middle-income countries (Armenia, Macedonia and Georgia) are among the countries with the lowest levels of manual tasks, comparable to those in high-income countries such as Belgium and France. These differences should be interpreted with caution. In all further analyses, we will focus on the other three measures and the RTI.

Turning into routine task intensity which measures the relative role of routine and non-routine tasks, we find that the cross-country differences in RTI are noticeably larger according to the country specific measures (standard deviation 0.27) than according to 0*NET (standard deviation 0.19) which assumes that occupations are identical around the world). Moreover, the relationship between countries' RTI and development level is stronger in the case of country-specific RTI (Figure 5).

Importantly, our survey-based measures show large cross-country differences in the relative routine-intensity of tasks in particular occupations (Figure 6). Among workers in high-skill occupations (ISC01 – managers, ISC0 2 – professionals, ISC0 3 – technicians) individuals in the more developed countries consistently perform less routine-intensive tasks than those in poorer countries. For those in middle- or low-skill occupations, the relationship between GDP per capita and relative routine intensity is inconsistent. Among sales and services workers (ISC0 5), workers in richer countries do less routine work. However, among craft and related trades workers (ISC0 7), and plant and machine operators (ISC08), there is an inverse-U shaped relationship between routine task intensity and development level of a country. Finally, among clerical workers (ISC0 4) and workers in elementary occupations (ISC09), the cross-country differences are highly variable but are not correlated with the level of GDP per capita.¹⁹

¹⁹ The standard deviation of country-specific average RTI is comparable for various occupations: it ranges from 0.22 (elementary occupations) to 0.28 (managers).

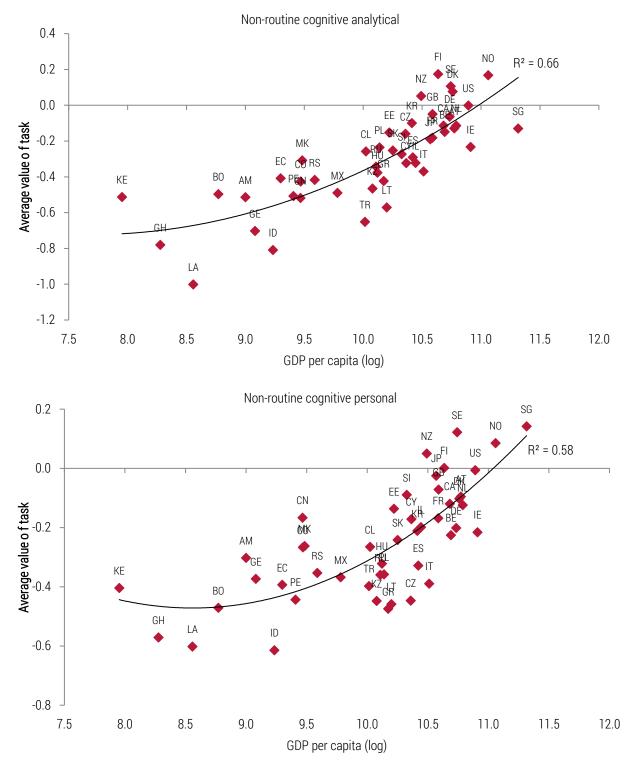


Figure 4. The average values of tasks by countries' development level (GDP per capita).

Note: for each task content, the 0 is set at the US average value and 1 corresponds to one standard deviation of this particular task content value in the US. GDP per capita in PPP, current international \$, country averages for 2011-2016. Source: own calculations based on PIAAC, STEP, CULS (tasks), and World Bank data (GDP).

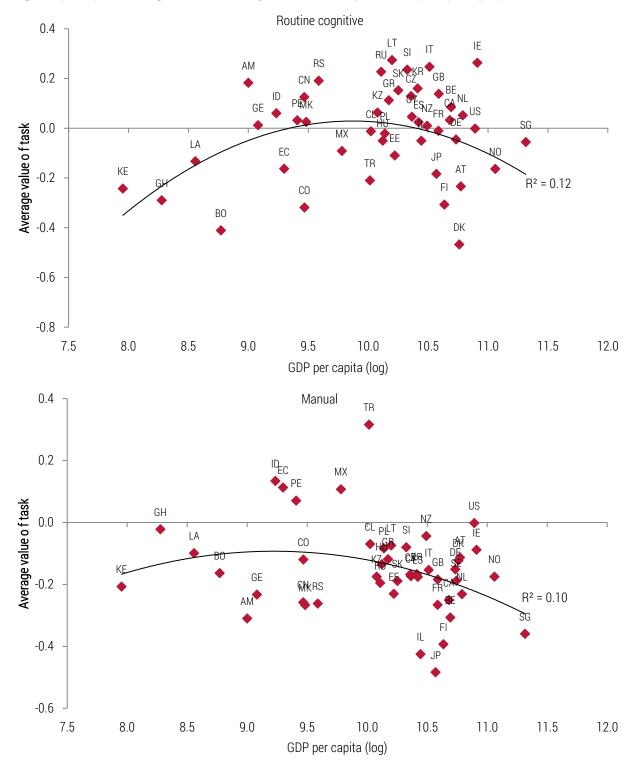
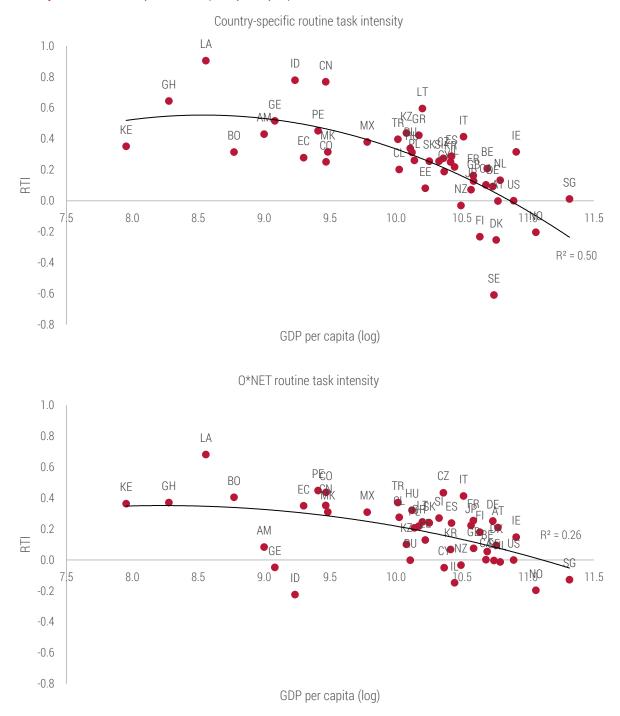
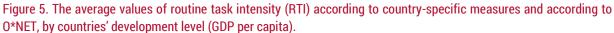


Figure 4 (cont'd). The average values of tasks by countries' development level (GDP per capita).

Note: for each task content, the 0 is set at the US average value and 1 corresponds to one standard deviation of this particular task content value in the US. GDP per capita in PPP, current international \$, country averages for 2011-2016. Source: own calculations based on PIAAC, STEP, CULS (tasks), and World Bank data (GDP).





Note: for each task content, the 0 is set at the US average value and 1 corresponds to one standard deviation of this particular task content value in the US. GDP per capita in PPP, current international \$, country averages for 2011-2016. Source: own calculations based on PIAAC, STEP, CULS (tasks), and World Bank data (GDP).

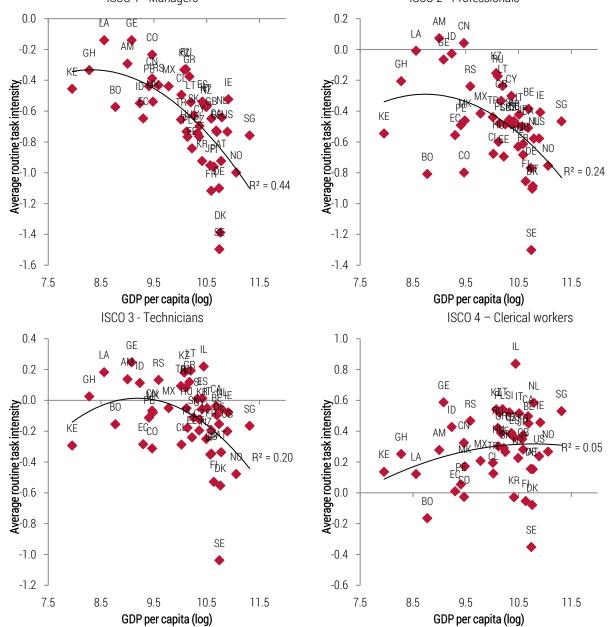
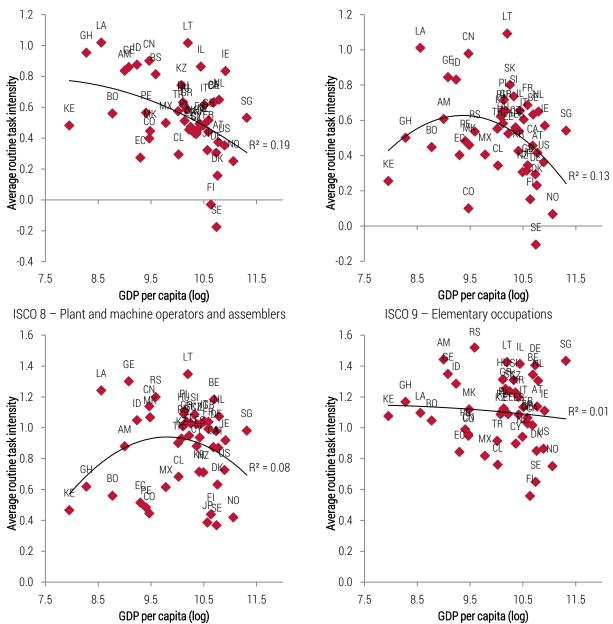


Figure 6. Average values of routine intensity of tasks (RTI) by 1-digit occupations by countries' development level. ISC0 1 - Managers ISC0 2 - Professionals





Note: the horizontal axis denotes GDP per capita, PPP (international \$, country averages for 2011-2016). We omit the occupational group ISCO 6 (Skilled agricultural workers) because of small sample sizes, especially in countries where surveys covered only urban areas.

Source: own calculations based on PIAAC, STEP, CULS, O*NET and World Bank data.

5. Determinants of Task Differences Across Countries

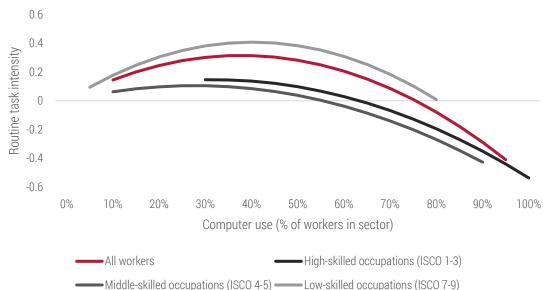
5.1 Determinants of task differences among all workers and by occupational groups

Regression results

We begin with discussing the results of a benchmark regression estimated for all workers, as well as for workers in high (ISCO 1-3), middle (ISCO 4-5) and low-skilled (ISCO 7-9) occupations, presented in Table 4 and Figure 7.

We find that better access to technology is associated with lower routine intensity of tasks performed by workers, especially in countries and sectors where more than 40% of workers use a computer. In countries and sectors with the share of computer use below 40%, the relationship between computer use and RTI is rather flat (Figure 7). We confirm this finding by re-estimating the baseline regressions with fixed effects for country-sector computer use share deciles instead of a continuous computer use share variable (Figure F1 in the Appendix). The estimated effects of computer use are sizeable. A 25 pp. higher share of computer use, equivalent to the difference between the US (75%) and China (50%), would translate into RTI being lower by 0.27 (of standard deviation of RTI in the US), which is equivalent to 40% of the difference between average RTI in the US and China. We also find evidence that the routine-replacing function of computers is stronger among workers in office or services jobs that usually require higher skills (ISCO 1-3 and ISCO 4-5) than among workers in low-skilled occupations (ISCO 7-9, Figure 7).

We also examine the impact of country level ICT stock and sector-level robot stock (both expressed in per worker terms) for a subsample of 32 countries with available data (Table F3 in Appendix). We find that higher level of ICT stock is associated with lower RTI: a level of ICT capital stock higher by 2.75 of standard deviation in our sample, which is equivalent to the difference between the US and China, is associated with RTI lower by 0.12, which is equivalent to 15% of a difference between the average worker in the US and the average worker in China. This effect is particularly large among workers in high-skilled occupations (Table F3). The relationship between the level of robot stock and RTI is also negative, but significant only for middle-skilled workers (Table F3).





Note: Based on the estimates presented in Table 4. For each category of workers, we select a range of computer use which includes 90% of workers in each category (we omit bottom and top 5%).

Source: own estimations based on PIAAC, STEP, CULS, World Bank and UIBE GVC Indicators data.

We find that globalization also plays an important role. In a country at the average GPD per capita level in our sample (e.g. Central Eastern European countries such as Slovakia or Estonia, or Chile), a higher foreign value added (FVA) share in domestic production is associated with a higher RTI. Thus, workers in sectors that specialize in smaller segments of global value chains (e.g., assemblers of final products) tend to perform more routine-intensive tasks. This effect is particularly strong for workers in low-skilled occupations (ISCO 7-9), followed by workers in middle-skilled occupations (ISCO 4-5), but is insignificant for workers in high-skilled occupations (Table 4). Importantly, the coefficient on the interaction term between FVA share and In(GDP per capita) is negative. This means that higher GDP attenuates the effect of FVA share on RTI: the effect becomes zero for a country with GDP equal to approx. 133% of the average in our sample, which is comparable to the US. At the same time, the positive impact of FVA share on RTI is nearly twice as great in countries with GDP per capita at about 25% of the mean in our sample, e.g. Georgia, Armenia, or Indonesia.

The second globalization measure, FDI share of GDP, is barely insignificant in the regression for all workers, but this masks heterogeneity across occupations. FDI is positively associated with RTI among workers in high-skill occupations, but negatively associated with RTI among workers in low-skill occupations. However, the magnitude of effects pertaining to FVA share is much larger than that of FDI. For instance, a 25 pp. higher FVA share, which is a difference between the US and countries most specialized in smaller segments of global value chains (e.g. small Central Eastern European countries) is associated with RTI in low-skilled occupations being higher by 0.16 (of the US standard deviation), which is equivalent to about 50% of the RTI difference between the US and these small CEE countries. But a 30 pp. difference in FDI share between the US and these CEE countries would translate into RTI being lower by only 0.005 of the US standard deviation. Overall, our results are consistent with theories arguing that routine jobs are easier to offshore and so poorer countries may specialize in them (Grossman and Rossi-Hansberg, 2008). Our findings are also robust to the choice of the GVC participation measure – the results of the estimation with forward linkage-based measure of participation in GVCs are presented in Table F2 in Appendix F. They are consistent with the baseline results obtained for the backward linkage-based measure (Table 4).

Next, we turn to the skill variables. Workers with higher education levels and higher literacy are more likely to perform fewer routine tasks, overall and within particular occupational groups. A worker with the highest literacy proficiency (level 4-5) is expected to perform 0.27 (of the US standard deviation) less routine-intensive task than an otherwise identical worker with a lower medium literacy proficiency (level 2). We also find that the workers performing more routine-intensive jobs are more likely to be female and young (aged 16-24). However, the relationship between age and RTI varies among occupational groups. In high-skilled occupations (ISCO 1-3), older individuals perform significantly less routine-intensive tasks, but in middle- (ISCO 4-5) and low-skilled occupations (ISCO 7-9), older workers perform more routine-intensive tasks, especially if aged over 55. This difference may suggest that experience and firm- or sector-specific knowledge can play a more important role for allocation of workers to tasks among high-skilled occupations than among middle- and low-skilled occupations.

The sector of employment also matters for the routine intensity of tasks performed by workers. On one hand, workers in service sectors, such as education; or arts, entertainment and recreation perform less routine-intensive jobs than workers in wholesale and retail trade, and repairs (the reference sector). On the other hand, workers in transportation and storage perform more routine-intensive tasks than workers in trade and repairs. However, there are notable heterogeneities of these sector effects across occupational groups. In many services which employ mainly workers in high-skilled occupations (e.g. information and communication, financial and insurance activities, education, and human health and social work activities), the minority of workers in low-skilled occupations, who

perform low-end support jobs, the intensity of routine tasks is significantly higher than among workers in low-skilled occupations in trade and repairs. Similarly, we find that workers in high-skilled occupations in real estate activities, financial and insurance activities, or health exhibit higher RTI than high-skilled workers in trade and repairs.

As an additional robustness check, we re-estimate our benchmark specification using averages of all variables at the country-sector level (Table F4 in Appendix F). The results show a negative relationship between the probability of computer use and skills, and RTI. The results for globalization variables are also close to those found in the pooled, worker-level regression. However, the coefficients pertaining to the employment shares of educational groups are not significant at the sector level. This suggests that the significance of education in the worker-level regressions reflects the allocation of less routine tasks to better educated workers within sectors.

	All workers	High-skilled occupations (ISCO 1-3)	Middle-skilled occupations (ISCO 4-5)	Low-skilled occupations (ISCO 7-9)
Computer use	1.671***	0.928***	0.762	2.029***
	(0.353)	(0.357)	(0.496)	(0.415)
Computer use ^2	-2.213***	-1.465***	-1.374***	-2.523***
	1.671***	0.928***	0.762	2.029***
Foreign Value Added (FVA)	0.235**	-0.142	0.354**	0.665***
share	(0.104)	(0.105)	(0.153)	(0.126)
Ln(GDP per capita) –	0.037	-0.019	0.061	0.075
mean(Ln(GDP per capita))	(0.042)	(0.043)	(0.061)	(0.050)
FVA share *	-0.172	-0.076	-0.194	0.034
[Ln(GDP pc) -mean(Ln(GDP pc)]	(0.106)	(0.118)	(0.173)	(0.122)
FDI / GDP	0.005	0.022***	0.003	-0.018***
	(0.005)	(0.006)	(0.008)	(0.007)
Education: primary	0.283***	0.145***	0.264***	0.159***
	(0.015)	(0.027)	(0.018)	(0.020)
Education: tertiary	-0.499***	-0.275***	-0.215***	-0.171***
	(0.016)	(0.016)	(0.020)	(0.033)
Literacy skills level:	0.099***	0.030	0.061**	0.089***
1 or lower	(0.015)	(0.023)	(0.024)	(0.021)
Literacy skills level: 3	-0.132***	-0.092***	-0.050**	-0.046**
	(0.012)	(0.014)	(0.019)	(0.022)
Literacy skills level:	-0.271***	-0.191***	-0.043	-0.160***
4 and 5	(0.018)	(0.018)	(0.027)	(0.042)
Female	0.240***	0.225***	0.197***	0.338***
	(0.011)	(0.012)	(0.017)	(0.023)
Age: 16-24	0.203***	0.211***	0.184***	0.124***
	(0.016)	(0.025)	(0.023)	(0.020)
Age: 35-44	-0.056***	-0.057***	-0.027*	-0.056***
	(0.010)	(0.013)	(0.016)	(0.019)
Age: 45-54	-0.021*	-0.058***	0.008	0.011
	(0.012)	(0.014)	(0.020)	(0.019)

Table 4. The correlates of routine task intensity (RTI) at the worker level

Age: 55-65	0.022	-0.041**	0.097***	0.061***
	(0.015)	(0.018)	(0.023)	(0.022)
Agriculture [A]	0.073	-0.052	-0.209	0.079
Agriculture [A]	(0.066)	(0.084)	(0.127)	(0.079)
Mining [D]	-0.056	-0.045	-0.213*	-0.062
Mining [B]	(0.073)	(0.082)	(0.117)	(0.115)
Manufacturing [0]	0.006	-0.063	-0.194***	-0.072
Manufacturing [C]	(0.049)	(0.049)	(0.062)	(0.055)
Electricity & Water supply	0.084	0.059	-0.075	0.190***
[D+E]	(0.054)	(0.057)	(0.125)	(0.055)
	-0.099*	-0.147***	-0.203***	-0.184***
Construction [F]	(0.051)	(0.052)	(0.076)	(0.059)
Transportation and storage	0.155***	-0.076	-0.010	0.075
' [H]	(0.049)	(0.052)	(0.078)	(0.058)
Accommodation and food	-0.016	-0.128*	-0.036	0.039
service [I]	(0.056)	(0.066)	(0.069)	(0.072)
Information and	0.000	0.073	0.154*	0.377***
communication [J]	(0.069)	(0.080)	(0.084)	(0.094)
	0.258***	0.341***	0.192**	0.816***
Financial and insurance [K]	(0.073)	(0.075)	(0.092)	(0.135)
Real estate & Professional	0.081	0.149*	0.159*	0.150
[L]	(0.063)	(0.081)	(0.082)	(0.118)
	-0.028	0.001	0.046	0.242***
Administrative [M+N]	(0.051)	(0.055)	(0.058)	(0.062)
	0.057	0.140**	0.044	0.312***
Public administration [0]	(0.058)	(0.056)	(0.073)	(0.069)
	-0.205***	-0.055	0.028	0.304***
Education [P]	(0.051)	(0.049)	(0.085)	(0.065)
	-0.004	0.217***	0.056	0.286***
Human health [Q]	(0.047)	(0.044)	(0.058)	(0.077)
	-0.235***	-0.100*	-0.039	-0.001
Arts [R]	(0.058)	(0.057)	(0.056)	(0.096)
	-0.261***	-0.214***	-0.300***	-0.083
Other service [S]	(0.056)	(0.055)	(0.069)	(0.068)
	0.349***	-0.003	0.036	0.263**
Activities of household [T]	(0.101)	(0.366)	(0.119)	(0.124)
Extraterritorial organizations	0.150	0.187	-0.154	0.472
[U]	(0.108)	(0.120)	(0.246)	(0.348)
No. of observations	166,542	68,027	52,906	45,609
R-squared	0.222	0.115	0.088	0.084

Note: *** p<0.01, ** p<0.05, * p<0.1. Standard errors in parentheses. We use standardized weights that give each country equal weight. The reference levels are: age 25-34, secondary education, wholesale and retail trade; repair of motor vehicles and motorcycles (ISIC G), lower medium literacy skills (level 2). The coefficients for interactions between sector fixed effects and Ln(GDP per capita) are presented in Table F1 in Appendix F.

Source: own estimations based on PIAAC, STEP, CULS World Bank, and UIBE GVC Indicators data.

Decompositions of cross-country differences

Our model accounts for more than 50% of the cross-country variance in RTI among all workers and among workers in high-skilled occupations, and for 25%-30% of the variance among workers in middle- and low-skilled occupations (Table 5). In each case, most of the variance is attributed to technology (computer use): about 40% in the pooled sample and among workers in high-skilled, and 30% and 25% among workers in middle- and low-skilled occupations, respectively. Supply of skills is the factor with second most relevant contribution overall (28%) and among workers in high-skilled occupations (10%). However, among workers in middle- and low-skilled occupations, globalization is the factor with the second largest contribution (7%), and the contribution of supply of skills is small.²⁰ The fact that skills account for much more of the cross-country variance when looking at all workers compared to variation within particular occupational skill groups suggests that a main influence of skills is its effect on occupational structure (we explore this further in next subsection).²¹ Finally, the contribution of structural change for all workers and for workers in high and for workers in middle-skilled occupations is negative. It's because the employment shares of some typically non-routine sectors (e.g. manufacturing) are in some low- and middle-income countries are lower than in the high-income countries.²²

	Technology	Globalization	Structural Change	Supply of skills	Total
All workers	41.3	12.4	-23.2	27.6	58.1
High-skilled occupations (ISCO 1-3)	39.6	6.5	-6.5	10.5	50.1
Middle-skilled occupations (ISCO 4-5)	29.3	7.3	-17.3	4.4	23.6
Low-skilled occupations (ISCO 7-9)	24.8	7.3	0.3	0.6	33.1

Table 5. Decomposition of cross-countr	v variance of RTI by fundamen	tal factors (% of total variance)
Table 5. Decomposition of cross-count	y variance of hit by fundamen	(a) Iduluis, (% UI lutal valiance)

Note: the contributions of particular factors to RTI variance, σ_k , calculated in line with equation (4) using the model presented in Table 4. Source: own estimations based on PIAAC, STEP, CULS, World Bank and UIBE GVC Indicators data.

Next, we decompose the gaps between average RTI in different countries and the US, which we take as the benchmark. We group countries into three types: low- and middle-income countries, the bottom high-income countries (including those in Southern, Central and Eastern Europe, as well as Chile and South Korea), and top high-income countries (mainly in North America, Western Europe, and Australasia, plus Singapore, Japan, and Israel). Decomposition results for gaps between every country and the US are reported in Appendix H.

²⁰ These results hold if we control for more technology variables (robot and ICT stocks per worker) and calculate decompositions based on regression results presented in Table F2 in Appendix F1. These results are available upon request.

²¹ We can conduct a similar variance decomposition analysis for individual-level RTI. We find that explained variance is much smaller (20% for all workers, 8% to 12% for different occupation skill groups). Skills account for the lion's share of explained variance, which is expected given that it is the only category for which we employ individual data. The relative importance of technology, globalization, and structural change are similar to the results for country differences.

²² The within-sector differences in RTI between less and more developed countries are substantial but this effect is of course attributed to other factors.

	Low- and Middle- Income Countries	Bottom High- Income Countries	Top High-Income Countries	US
RTI	0.45	0.28	0.01	0.00
Computer use	0.36	0.59	0.76	0.75
Log of GDP per capita (demeaned)	-1.23	0.15	1.07	1.29
FDI stock/GDP	0.44	1.19	0.79	0.35
FVA Share	0.14	0.25	0.19	0.08
Education: primary	0.34	0.17	0.15	0.10
Education: secondary	0.34	0.49	0.43	0.48
Education: tertiary	0.32	0.34	0.42	0.42
iteracy skills level: 1 or lower	0.49	0.18	0.13	0.14
Literacy skills level: 2	0.35	0.38	0.31	0.31
Literacy skills level: 3	0.14	0.37	0.41	0.40
Literacy skills level: 4 and 5	0.02	0.08	0.15	0.15
Female	0.45	0.46	0.48	0.49
Age: 16-24	0.17	0.08	0.12	0.15
Age: 25-34	0.29	0.25	0.22	0.23
Age: 35-44	0.24	0.28	0.25	0.22
Age: 45-54	0.19	0.25	0.25	0.23
Age: 55-65	0.10	0.14	0.16	0.18
Agriculture [A]	0.015	0.019	0.008	0.008
Mining [B]	0.008	0.006	0.004	0.005
Manufacturing [C]	0.167	0.193	0.140	0.112
Electricity & Water supply [D+E]	0.013	0.018	0.013	0.010
Construction [F]	0.072	0.084	0.069	0.066
Trade and repairs [G]	0.221	0.152	0.138	0.117
Transportation and storage [H]	0.065	0.058	0.054	0.043
Accommodation and food service [I]	0.062	0.055	0.050	0.072
Information and communication [J]	0.024	0.030	0.040	0.043
Financial and insurance [K]	0.020	0.029	0.037	0.048
Real estate & Professional [L]	0.005	0.008	0.011	0.015
Administrative [M+N]	0.072	0.078	0.098	0.101
Public administration [0]	0.043	0.066	0.059	0.060
Education [P]	0.085	0.088	0.088	0.090
Human health [Q]	0.042	0.063	0.140	0.137
Arts [R]	0.016	0.018	0.021	0.026
Other service [S]	0.045	0.024	0.025	0.032
Activities of household [T]	0.025	0.011	0.004	0.015
Extraterritorial organizations [U]	0.001	0.001	0.000	0.000

Table 6. Average levels of RTI and explanatory variables by country groups

Source: own calculations based on PIAAC, STEP, CULS World Bank, and UIBE GVC Indicators data.

The average RTI scores are much higher for low- and middle-income countries (0.45 US standard deviations, Table 6) and bottom high-income countries (0.28) than for the top high-income group and the US (0.01 and 0). Low- and middle-income countries exhibit much lower shares of workers who use computer (36%) compared to the top high-income countries (76%) and the US (75%). In terms of skill, the low- and middle-income countries have fewer older workers and double the share of workers with education level of primary school and below. Notably, 49% of workers in low-and middle-income countries are at the lowest literacy level, compared to just 13% and 14% in top high-income countries and the US; and just 14% of workers in low-and middle-income countries are at the at least upper medium literacy level (3-5), compared to 45% in the bottom high-income countries, and 56% and 55% in top high-income countries and the US, respectively. Finally, specialization in global value chains as captured by foreign value added share is highest in the bottom high income group (0.25) compared to just 0.14 in low- and middle-income countries, 0.19 in top high income countries, and 0.08 in the US. The gap in GDP per capita is about 2.5 points in log scale, implying that GDP per capita is more than 250% greater in the top high-income countries compared to the low- and middle-income countries.

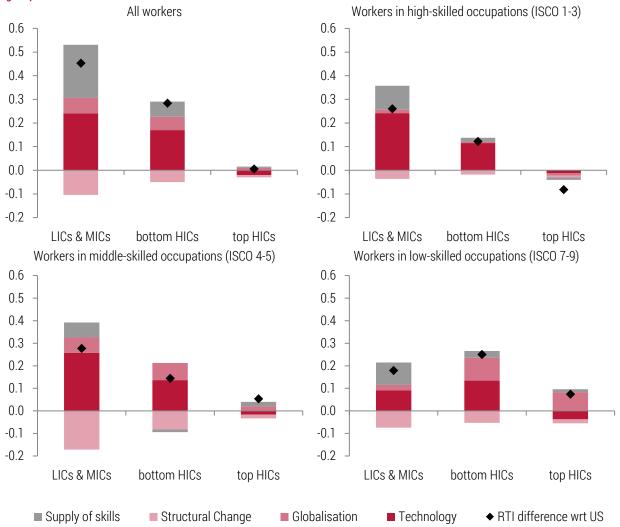


Figure 8. Regression-based decomposition of differences in RTI between particular countries and the US, by country groups.

Note: Results of decomposition (5) based on the estimates presented in Table 4, and averaged for country groups defined in Table 3. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US. Source: own estimations based on PIAAC, STEP, CULS, World Bank and UIBE GVC Indicators data.

The decomposition reveals that technology has the largest contribution to the RTI difference in all country groups, but the role of other factors differs between them (Figure 8). In low- and middle-income countries, the contribution of skill supply – lower in those countries than in the US – is nearly as large as that of technology. For the bottom high-income countries, the contribution of skill supply is much lower than in the low- and middle-income countries, and consequently, the RTI gaps with the US are about half as large on average compared to low- and middle-income countries. In both groups of countries, globalization plays a moderate role, with structural differences contributing negatively. For the top high-income countries, RTI gaps with the US are negligible.

Next, we conduct that decomposition separately for high-, middle-, and low-skilled occupations. Compared to the results for all workers, for those in high-skilled occupations, RTI gaps with the US are slightly smaller and technology explains even a larger share of these gaps (Figure 8). For low- and middle-income countries, a noticeable share of the RTI difference can be attributed to lower supply of skills among workers high-skilled occupations in these countries. For middle-skilled occupations, technology is the most important factor, followed by globalization. Skills play a role only for low- and middle-income countries. The relatively greater importance of technology to differences in RTI of high skill occupations is consistent with technology being complementary to non-routine cognitive tasks.

Finally, in low-skilled occupations, gaps with the US in other high-income countries are greater than for high- and middle-skilled occupations. Globalization is by far the most important factor in accounting for these gaps, followed by technology. This suggests that even high-income countries, the division of labor between countries more and less narrowly specialized in GVCs is associated with noticeable differences in routine intensity of work in occupations such as assemblers, or plant and machine operators. In low- and middle-income countries, however, differences with the US among in low-skilled occupations are smaller than in other occupational groups, and can be mainly attributed to lower supply of skills, followed by technology.

5.2 Assessing the role of occupations

The results thus far have shown that most of the cross-country differences in job tasks are associated with differences in technology, supply of skills, and globalization. Having individual survey data on job tasks enables us to study the correlates of RTI without making any assumptions about the nature of work in different occupations. Nonetheless, given that much research on the nature of work focuses on occupations, it is of interest to investigate how much of the above-documented relationship between routine task intensity and the four fundamental forces is explained by differences in occupational structure and how much is due to differences in tasks within occupation. To examine this question, we estimate regression (3) which expands the benchmark specification with occupation fixed effects, τ_o . In this specification, the coefficients on the variables for the four main factors capture their influence among workers in the same occupation. Thus, by comparing the coefficients with those estimated using the baseline specification without occupation fixed effects, we can infer how much of the relationship between the four factors is captured by their impact on occupational structure and how much is a within-occupation association.

The coefficients on the occupation dummies are in line with intuition: workers in high-skilled occupations (ISCO 1-3) perform less routine-intensive tasks than clerical workers (ISCO 4), while workers in low-skilled occupations (ISCO 7-9) and sales and services workers (ISCO 5) perform more routine-intensive tasks (Table 7). However, the four fundamental factors still strongly predict differences in routine intensity even after controlling for occupations. Although the absolute sizes of the coefficients pertaining to education and literacy are somewhat smaller than in the benchmark specification (Table 4), none of them loses statistical significance. The coefficients pertaining to computer use and globalization variables change little and remain significant.

	All workers	High-skilled occupations (ISCO 1-3)	Middle-skilled occupations (ISCO 4-5)	Low-skilled occupations (ISCO 7-9)
Computer use	1.439***	0.900**	0.755	1.967***
	(0.348)	(0.372)	(0.489)	(0.414)
Computer use ^2	-1.888***	-1.469***	-1.373***	-2.360***
	1.439***	0.900**	0.755	1.967***
Faraign Value Added above	0.325***	-0.120	0.355**	0.673***
Foreign Value Added share	(0.100)	(0.110)	(0.149)	(0.130)
Ln(GDP per capita) –	0.008	-0.028	0.064	0.026
mean(Ln(GDP per capita))	(0.040)	(0.048)	(0.060)	(0.056)
FVA share *	-0.112	-0.044	-0.184	0.065
Ln(GDP pc) -mean(Ln(GDP pc)]	(0.106)	(0.124)	(0.171)	(0.140)
FDI / GDP	0.004	0.021***	0.004	-0.016***
	(0.005)	(0.006)	(0.008)	(0.006)
Education: primary	0.180***	0.153***	0.257***	0.134***
	(0.014)	(0.027)	(0.018)	(0.019)
Education: tertiary	-0.204***	-0.215***	-0.205***	-0.146***
	(0.014)	(0.016)	(0.019)	(0.033)
Literacy skills level:	0.055***	0.029	0.056**	0.081***
1 or lower	(0.015)	(0.023)	(0.024)	(0.021)
Litere ev elville level, 2	-0.062***	-0.083***	-0.044**	-0.036
Literacy skills level: 3	(0.011)	(0.014)	(0.020)	(0.022)
Literacy skills level:	-0.140***	-0.168***	-0.036	-0.140***
4 and 5	(0.015)	(0.018)	(0.027)	(0.040)
Female	0.226***	0.203***	0.207***	0.275***
	(0.011)	(0.011)	(0.017)	(0.023)
Age: 16-24	0.156***	0.191***	0.182***	0.100***
	(0.014)	(0.025)	(0.023)	(0.020)
Age: 35-44	-0.035***	-0.034***	-0.028*	-0.046**
	(0.009)	(0.012)	(0.016)	(0.018)
Age: 45-54	0.001	-0.027**	0.008	0.021
5	(0.011)	(0.014)	(0.020)	(0.018)
Age: 55-65	0.053***	-0.005	0.098***	0.073***
-	(0.012)	(0.018)	(0.023)	(0.021)
ISCO 1	-0.766***	-	-	-
	(0.021)			
ISCO 2	-0.576***	0.191***	-	-
	(0.020)	(0.017)		
ISCO 3	-0.344***	0.413***	-	-
	(0.018)	(0.016)		

Table 7. The correlates of routine task intensity (RTI) at the worker level, including occupations

ISCO 5	0.129***	-	0.128***	-
	(0.020)		(0.021)	
ISCO 7	0.210***	-	-	-0.416***
	(0.022)			(0.024)
ISCO 8	0.528***	-	-	-0.099***
	(0.023)			(0.023)
ISCO 9	0.593***	-	-	-
	(0.022)			
Sector fixed effects	Yes	Yes	Yes	Yes
Observations	166,542	68,027	52,906	45,609
R-squared	0.314	0.139	0.090	0.111
· · · · · · · · · · · · · · · · · · ·				

Note: *** p<0.01, ** p<0.05, * p<0.1. Standard errors in parentheses. We use standardized weights that give each country equal weight. The reference levels are: age 25-34, secondary education, wholesale and retail trade; repair of motor vehicles and motorcycles (ISIC G), lower medium literacy skills (level 2). Clerical support workers (ISCO 4) are the reference group in the regressions for all workers, for middle-skilled occupations and for offshorable and non-offshorable occupations. Managers (ISCO 1) and Elementary occupations (ISCO 9) are the reference groups in regressions for high-skilled and low-skilled occupations, respectively.

Source: own estimations based on PIAAC, STEP, CULS, World Bank, and UIBE GVC Indicators data.

Next, we conduct the cross-country variance decomposition and decomposition of gaps with the US adding the occupational structure as an additional factor. We find that occupations have a noticeable contribution to the cross-country variation in RTI for all workers (22%). However, the total variance explained by the model (57.5%) is virtually the same as in the specification with no occupational fixed effects (58.1%, Table 8). The contributions attributed to other factors, especially to the supply of skills, are lower than for the benchmark specification. Still, the contribution of technology remains larger than the contribution of occupations. When we analyze the importance of occupations separately for high-, middle-, and low-skilled occupation groups (not reported here and available upon request), we find that the occupation dummies have very little explanatory power. This suggests that only differences in the shares of broad occupation categories are meaningful for explaining task content differences across countries, and that they are to a large extent related to cross-country differences in skill supply.

variance)						
	Technology	Globalization	Structural Change	Supply of skills	Occupations	Total
With occupation fixed effects	34.4	8.5	-20.1	12.9	21.8	57.5
Without occupation fixed effects	41.3	12.4	-23.2	27.6		58.1

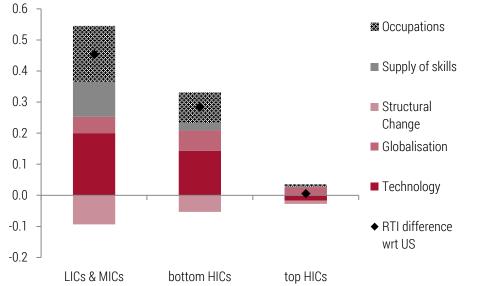
Table 8. Decomposition of cross-country variance of RTI by fundamental factors, controlling for occupations (% of total variance)

Note: the contributions of particular factors to RTI variance, σ_k , calculated in line with equation (4) using the model presented in Table 7. Source: own estimations based on PIAAC, STEP, CULS, World Bank and UIBE GVC Indicators data.

In explaining gaps in RTI between low- and middle-income countries and the US, we find that occupational structure contributes only about one fifth of the explained gap, with four fifths of the gap explained by the association of fundamental factors and within-occupation differences in RTI across countries (Figure 9).

Overall, we find, perhaps surprisingly, that occupations are of limited importance in explaining cross-country differences in routine task intensity. Their contribution amounts to 40% of the total difference for low- and middle-income countries, and 35% for bottom high-income countries (Figure 9). Most of the association between fundamental factors and RTI appear in differences within occupation groups. This highlights the importance of collecting survey-based task data to understand the nature of work in specific countries.

Figure 9. Regression-based decomposition of differences in RTI between particular countries and the US, controlling for occupations, by country groups.



Note: Results of decomposition (5) based on the estimates presented in Table 9, and averaged for country groups defined as in Table 4. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US. Source: own estimations based on PIAAC, STEP, CULS, World Bank and UIBE GVC Indicators data.

6. Summary and conclusions

We have developed a novel dataset that measures the task content of jobs at the individual worker level for a large number of countries at different stages of economic development. The new survey-based measures are validated to be consistent with US O*NET-based task content measures that have been widely used in the existing literature on job tasks. A key advantage of the new measures is that they can distinguish between differences in task content among workers who have the same occupation but live in different country environments.

Our results show that there are substantial cross-country differences in the routine-intensity of job tasks, both at the national level and within specific occupations. The differences in tasks across countries at different stages of development are much greater than could be explained by differences in occupational structure. Not surprisingly, jobs in the most developed countries involve the most non-routine cognitive analytical and non-routine cognitive interpersonal tasks, and often have the least manual tasks, while the opposite is true for developing and emerging economies. Routine cognitive tasks are lowest in the least and most developed countries, and highest in Eastern and Southern European countries, suggesting an inverse U-shaped relationship between the role of routine cognitive work and development level. Moreover, cross-country differences in routine task intensity are most strongly related to the differences in GDP per capita for high-skilled occupations, with no systematic correlation for middle- and low-skill occupations.

We have estimated a regression that captures the association between the relative routine task intensity (RTI) of jobs and four fundamental forces: technology, globalization, structural change, and supply of skills. We have used these results to decompose the extent to which cross-country differences in relative routine task intensity are statistically associated with these different factors, both in terms of cross-country variance in mean RTI and RTI gaps between the US and groups of countries sorted by GDP per capita. Consistent with much recent literature emphasizing the influence of technology on the nature of work, we find that technology plays the largest role in

explaining cross-country differences in RTI, followed by skills and globalization. Sector structures have the least explanatory power. However, we have also found interesting heterogeneities in the impact of these factors for different types of occupations. Technology matters the most for high-skilled occupations, consistent with the complementarity between technology and non-routine cognitive tasks, while globalization matters the most for low-skill occupations which are more likely to involve routine tasks that are more easily outsourced from rich countries to poor countries.

Our work stresses the need to quantify the country-specific task content of jobs and identify differences between occupational task content in countries at different stages of development. It paves the way for a more comprehensive research on the distribution of tasks around the world that can account for the within-occupation and between-country variation in task demand.

References

- Acemoglu, D., Autor, D., 2011. Skills, Tasks and Technologies: Implications for Employment and Earnings, in: Card, D., Ashenfelter, O. (Eds.), Handbook of Labor Economics. Elsevier, pp. 1043–1171. https://doi.org/10.1016/S0169-7218(11)02410-5
- Akerman, A., Gaarder, I., Mogstad, M., 2015. The Skill Complementarity of Broadband Internet. The Quarterly Journal of Economics 130, 1781–1824. https://doi.org/10.1093/qje/qjv028
- Arias, O.S., Sánchez-Páramo, C., Dávalos, M.E., Santos, I., Tiongson, E.R., Gruen, C., de Andrade Falcão, N., Saiovici, G., Cancho, C.A., 2014. Back to Work: Growing with Jobs in Europe and Central Asia | Europe and Central Asia Reports. https://doi.org/10.1596/978-0-8213-9910-1
- Arntz, M., Gregory, T., Zierahn, U., 2017. Revisiting the risk of automation. Economics Letters 159, 157–160. https://doi.org/10.1016/j.econlet.2017.07.001
- Autor, D., Dorn, D., 2009. This Job Is "Getting Old": Measuring Changes in Job Opportunities Using Occupational Age Structure. American Economic Review 99, 45–51. https://doi.org/10.1257/aer.99.2.45
- Autor, D.H., 2013. The "task approach" to labor markets: an overview. J Labour Market Res 46, 185–199. https://doi.org/10.1007/s12651-013-0128-z
- Autor, D.H., Dorn, D., 2013. The Growth of Low-Skill Service Jobs and the Polarization of the US Labor Market. American Economic Review 103, 1553–1597. https://doi.org/10.1257/aer.103.5.1553
- Autor, D.H., Handel, M.J., 2013. Putting Tasks to the Test: Human Capital, Job Tasks, and Wages. Journal of Labor Economics 31, 59–96.
- Autor, D.H., Katz, L.F., Kearney, M.S., 2006. The Polarization of the U.S. Labor Market. American Economic Review 96, 189–194. https://doi.org/10.1257/000282806777212620
- Autor, D.H., Levy, F., Murnane, R.J., 2003. The Skill Content of Recent Technological Change: An Empirical Exploration. Q J Econ 118, 1279–1333. https://doi.org/10.1162/003355303322552801
- Bárány, Z.L., Siegel, C., 2018. Job Polarization and Structural Change. American Economic Journal: Macroeconomics 10, 57–89. https://doi.org/10.1257/mac.20150258
- de la Rica, S., Gortazar, L., Lewandowski, P., 2020. Job Tasks and Wages in Developed Countries: Evidence from PIAAC. Labour Economics 65, 101845. https://doi.org/10.1016/j.labeco.2020.101845
- Deming, D.J., 2017. The Growing Importance of Social Skills in the Labor Market. Q J Econ 132, 1593–1640. https://doi.org/10.1093/qje/qjx022
- Du, Y., Jia, P., Cheng, J., 2017. Structural Changes in Labor Market, Job Tasks and Skill Demand. Studies in Labor Economics 5, 30–49.
- Du, Y., Park, A., 2018. Changing Demand for Tasks and Skills in China, Background Report for the World Bank Group (WBG)-Development Research Center under the State Council (DRC) report on New Drivers of Growth in China.
- Eden, M., Gaggl, P., 2020. Do Poor Countries Really Need More IT? World Bank Econ Rev 34, 48–62. https://doi.org/10.1093/wber/lhy022
- Firpo, S., Fortin, N.M., Lemieux, T., 2011. Occupational Tasks and Changes in the Wage Structure (No. 5542), IZA Discussion Papers. Institute of Labor Economics (IZA).
- Frey, C.B., Osborne, M.A., 2017. The future of employment: How susceptible are jobs to computerisation? Technological Forecasting and Social Change 114, 254–280. https://doi.org/10.1016/j.techfore.2016.08.019
- Goos, M., Manning, A., Salomons, A., 2014. Explaining Job Polarization: Routine-Biased Technological Change and Offshoring. American Economic Review 104, 2509–2526. https://doi.org/10.1257/aer.104.8.2509

- Grossman, G.M., Rossi-Hansberg, E., 2008. Trading Tasks: A Simple Theory of Offshoring. American Economic Review 98, 1978–1997. https://doi.org/10.1257/aer.98.5.1978
- Hardy, W., Keister, R., Lewandowski, P., 2018. Educational upgrading, structural change and the task composition of jobs in Europe. Economics of Transition and Institutional Change 26, 201–231. https://doi.org/10.1111/ecot.12145
- Hsieh, C.-T., Klenow, P.J., 2010. Development Accounting. American Economic Journal: Macroeconomics 2, 207–223. https://doi.org/10.1257/mac.2.1.207
- Hummels, D., Munch, J.R., Xiang, C., 2018. Offshoring and Labor Markets. Journal of Economic Literature 56, 981–1028. https://doi.org/10.1257/jel.20161150
- Jensen, J.B., Kletzer, L.G., 2010. Measuring Tradable Services and the Task Content of Offshorable Services Jobs, NBER book chapter series. National Bureau of Economic Research, Cambridge, Mass.
- Lewandowski, P., Keister, R., Hardy, W., Górka, S., 2020. Ageing of routine jobs in Europe. Economic Systems 44, 100816. https://doi.org/10.1016/j.ecosys.2020.100816
- Lo Bello, S., Sanchez Puerta, M.L., Winkler, H., 2019. From Ghana to America: The Skill Content of Jobs and Economic Development (No. 12259), IZA Discussion Papers. Institute of Labor Economics (IZA).
- Marcolin, L., Miroudot, S., Squicciarini, M., 2019. To be (routine) or not to be (routine), that is the question: a cross-country task-based answer. Ind Corp Change 28, 477–501. https://doi.org/10.1093/icc/dty020
- McKinsey, 2017. Future that works: AI, automation, employment, and productivity. McKinsey Global Institute.
- Michaels, G., Natraj, A., Van Reenen, J., 2014. Has ICT Polarized Skill Demand? Evidence from Eleven Countries over Twenty-Five Years. The Review of Economics and Statistics 96, 60–77. https://doi.org/10.1162/REST_a_00366
- Montresor, G., 2019. Job polarization and labour supply changes in the UK. Labour Economics 58, 187–203. https://doi.org/10.1016/j.labeco.2018.05.009
- Morduch, J., Sicular, T., 2002. Rethinking Inequality Decomposition, with Evidence from Rural China. The Economic Journal 112, 93–106. https://doi.org/10.1111/1468-0297.0j674
- Oldenski, L., 2012. Export Versus FDI and the Communication of Complex Information. Journal of International Economics 87, 312–322. https://doi.org/10.1016/j.jinteco.2011.12.012
- PIAAC Literacy Expert Group, 2009. PIAAC Literacy: A Conceptual Framework (No. 34), ECD Education Working Papers. OECD Publishing.
- Reijnders, L.S.M., de Vries, G.J., 2018. Technology, offshoring and the rise of non-routine jobs. Journal of Development Economics 135, 412–432. https://doi.org/10.1016/j.jdeveco.2018.08.009
- Rubin, D.B., 1987. Multiple Imputation for Nonresponse in Surveys. John Wiley and Sons Inc, New York City.
- Salvatori, A., 2018. The anatomy of job polarisation in the UK. Journal for Labour Market Research 52, 8. https://doi.org/10.1186/s12651-018-0242-z
- Spitz-Oener, A., 2006. Technical Change, Job Tasks, and Rising Educational Demands: Looking outside the Wage Structure. Journal of Labor Economics 24, 235–270. https://doi.org/10.1086/499972
- Wang, Z., Wei, S.-J., Yu, X., Zhu, K., 2017. Measures of Participation in Global Value Chains and Global Business Cycles, NBER working paper series. National Bureau of Economic Research, Cambridge, Mass.

World Bank, 2019. World Development Report 2019: The Changing Nature of Work. World Bank, Washington, DC.

Appendices

Appendix A. List of countries in PIAAC, STEP and CULS

PIAAC surveys include publicly available data representative of 37 countries. 23 in Round I: Austria, Belgium (Flanders), Canada, Cyprus (the area under the effective control of the Government of the Republic of Cyprus), Czechia, Denmark, Estonia, Finland, France, Germany, Ireland, Italy, Japan, South Korea, Netherlands, Norway, Poland, Russia (w/o Moscow municipal area), Slovakia, Spain, Sweden, UK (England and Northern Ireland), United States. 9 in Round II: Chile, Greece, Indonesia (Jakarta), Israel, Lithuania, New Zealand, Singapore (only permanent residents), Slovenia and Turkey. 6 in Round III: Ecuador, Hungary, Kazakhstan, Mexico, Peru and an additional round for United States. Moreover, a dataset with a supplementary sample for the first round is available for the United States via the US National Center for Education Statistics (NCES). We use the initial Round I sample for the USA with the supplementary sample from NCES.

We use STEP surveys for 9 countries: Armenia, Bolivia (four main capital cities – La Paz, El Alto, Cochabamba and Santa Cruz de la Sierra), Colombia (13 main metropolitan areas), Georgia (w/o Abkhazia and South Ossetia), Ghana, Kenya, Lao PDR (both urban and rural areas), Macedonia, and Serbia.

The 3rd wave of the CULS survey includes data on individuals in six large cities in China: Guangzhou, Shanghai, and Fuzhou on the coast, Shenyang in the northeast, Xian in the northwest, and Wuhan in central China.

Appendix B. Construction of task content measures based on US PIAAC and US O*NET

To construct the reference task content measures proposed by Acemoglu and Autor (2011), we use the Occupational Information Network (O*NET) database which contains extensive information on the occupations in the US. We merge the O*NET data with the US PIAAC data using the occupational crosswalks prepared by the O*NET Resource Center, the U.S. Bureau of Labor Statistics and the National Crosswalk Service Center, and adapted to the ISCO classification of occupations by Hardy et al. (2018).²³ ISCO is used in PIAAC, and 3-digit or 4-digit codes are available in the US PIAAC.²⁴ We apply our procedure at each level.

To calculate the task content of occupations, we follow Acemoglu and Autor (2011). First, we standardize the values $t_{j^o,i}$ of each task item j^o in the set of O*NET task items J^o , using the means $(\bar{t}_{j^o}^{US})$ and standard deviations $(\delta_{j^o}^{US})$ in the US PIAAC:

$$\forall_i \forall_j o_{\in J} o \ t_{i,j}^{std} = \frac{t_{i,j} o - \overline{t}_{j,0}^{US}}{\delta_{i,0}^{US}}, \tag{B1}$$

whereby *i* is a worker-level observation in the US PIAAC data. The set of O*NET task items, J^o , is presented in Appendix B. Second, we construct four task content measures: non-routine cognitive analytical, non-routine cognitive interpersonal, routine cognitive, and manual. Each task content measure is calculated as a sum of constituent task items (Table B1), except for the manual measure which is the sum of all items that define routine

²³ See: www.ibs.org.pl/resources [accessed: 2017-05-04].

²⁴ The dataset with 3-digit ISCO codes is available for researchers from National Center for Education Statistics. The 4-digit ISCO codes are included in the restricted dataset at the American Institutes for Research who have kindly run our code.

and non-routine manual task content measures in Acemoglu and Autor (2011). Each of these sums is then standardized to have a mean 0 and standard deviation 1 in the US PIAAC sample. Using one measure of manual tasks is not a limitation because the correlation between the non-routine and routine manual tasks in the US PIAAC is very high (85% across 3-digit ISCO occupations and 88% across 2-digit occupations).²⁵

Task content measure (7)	Task items (J)
Non-routine cognitive analytical	Analysing data/information
	Thinking creatively
	Interpreting information for others
Non-routine cognitive interpersonal	Establishing and maintaining personal relationships
	Guiding, directing and motivating subordinates
	Coaching/developing others
Routine cognitive	The importance of repeating the same tasks
	The importance of being exact or accurate
	Structured vs. unstructured work
Routine manual	Pace determined by the speed of equipment
	Controlling machines and processes
	Spending time making repetitive motions
Non-routine manual physical	Operating vehicles, mechanized devices, or equipment
	Spending time using hands to handle, control or feel objects, tools or controls
	Manual dexterity
	Spatial orientation

Table B1. Set of O*NET items, J^o, used in Acemoglu and Autor (2011) task contents measures

Source: Own elaboration based on Acemoglu and Autor (2011).

PIAAC and STEP surveys provide data on the job tasks performed by workers. In first step, we identified the set of potential items, $J^P = \{J^{P,NRCA}, J^{P,NRCP}, J^{P,RC}, J^{P,M}\}$, that are available in both surveys in the identical or almost identical form, and which could potentially be used to derive particular task content measures (see Table B2, and Appendix C for the full wording of questions and allowed answers). We chose between three and eight potential items for particular task content measures, except the manual content for which only one item ("working physically") is available in both STEP and PIAAC. We reverse the values of three variables considered for the routine cognitive measure ("changing order of tasks", "solving problems", "giving presentations"), so the higher is the value, the less common is a given phenomenon. To ensure comparability between STEP and PIAAC data, we rescale the answers to achieve the same value ranges. In particular, for PIAAC questions with five possible answers, except for "changing order of tasks", "solving problems", and "giving presentations", we consider four variants of binary variables, based on cutoffs available in the original answers (see Appendix C for details).

Our selection of questions is based on the similarities between PIAAC / STEP items and the O*NET items, and attributes of particular type of work (Autor, 2013). It is also consistent with the selections of Lo Bello et al. (2019), de la Rica et al. (2020) or Marcolin et al. (2019), shown in Appendix D, with three distinctions. First, those studies used either STEP data or PIAAC data and could have used any questions available in a given survey. We combine these two datasets so we can use only questions present in both surveys. Second, due to data availability we construct a manual task measure while Lo Bello et al. (2019) were able to distinguish between non-routine and routine manual tasks on the basis of STEP data. Third, in those studies the task measures were constructed in an

²⁵ Studies on the US (Autor and Price, 2013) or European countries (Lewandowski et al., 2017) found that routine and non-routine manual tasks are also highly correlated over time and follow similar trends.

arbitrary way, while we search for subsets of questions and cutoffs which provide the best proxy for the O*NET tasks in the US. While we end up with a similar number of items per task, our approach allows selecting the task items on basis of objective criteria.

Formally, we consider every subset of different questions allowed for a particular task measure:

$$t \in T = \{x_{c_1}^{j_1}, \dots, x_{c_l}^{j_k}: 2 \le k \le r_R, j_1 \ne j_2 \ne \dots \ne j_k, j_i \in J^{P,R}\},$$
(B2)

Where r_R is the number of questions considered for particular task content $R \in \{NRCA, NRCP, RC, M\}$ (Table B2), l is the number of variable variants c available for each question (1 or 4), and x are the values. Note that we don't allow two variants of the same question in a given subset, and consider only subsets with at least two variables. The total number of subsets considered for particular task content measures is shown in Table B2.

In the next step, we adapt the Acemoglu and Autor (2011) methodology to the PIAAC items. We standardize the worker-level values $x_{c_l,i}^{j_k}$ using the means $(\bar{t}_{x_{c_l}^{j_k}}^{US})$ and standard deviations $(\delta_{x_{c_l}^{j_k}}^{US})$ in the US:

$$\forall_{i} \forall_{j_{c_{l}}^{k} \in J^{P,R}} x_{c_{l},i}^{j_{k},std} = \frac{x_{c_{l},i}^{j_{k}} - \bar{t}_{x_{c_{l}}^{J_{k}}}^{US}}{\frac{\delta^{US}}{x_{c_{l}}^{j_{k}}}}, \tag{B3}$$

For each subset, we sum these standardized values and standardize those sums again within the US dataset. Then, we calculate (weighted) averages of these subset-specific values at the level of 3-digit and 4-digit ISCO occupations. Finally, we calculate the correlations between these occupation-specific averages and the relevant O*NET-based task content measures across 3-digit and 4-digit ISCO occupations in the US.

	Non-routine cognitive analytical (J ^{P,NRCA})	Non-routine cognitive personal (J^{P,NRCP})	Routine cognitive $(J^{P,RC})$	Manual (J^{P,M})
	Solving problems (1)	Supervising (1)	Changing order of tasks -	Physical
	Reading bills (4)	Collaborating (1)	reversed (1)	tasks (1)
	Reading news (4)	Making speeches or giving	Reading bills (4)	
	Reading professional	presentations (4)	Filling forms (4)	
Task items	journals (4)		Calculating fractions (4)	
	Advanced math (4)		Solving problems -	
	Calculating prices (4)		reversed (1)	
	Calculating fractions (4)		Making speeches or giving	
	Programming (4)		presentations - reversed (4)	
No. of subsets	156 221	18	4 982	1

Table B2. PIAAC and STEP questions considered for the measurement of particular task content measures, with number of variable variants (in brackets)

Note: 1 and 4 identify variables for which we use original questions (1), or four variants of binary variables based on cutoffs available in original question (4). For each task content measure except the manual measure, we consider only combinations that include at least two questions. Last row shows the number of subsets of variables considered for given task content measure. Source: Own elaboration.

For each task content measure, we use the following criteria to select the best subset of PIAAC items:

 We consider five subsets with the highest correlations with the relevant O*NET-based measure at the 3-digit, or at the 4-digit level of ISCO.

- A particular subset can be preferred over a subset with the higher correlation at the 4-digit level only if it has a higher correlation at the 3-digit level.
- The reversed version of variables used in the measure of routine cognitive tasks should use the same cutoffs as the original variables used in the measures of non-routine cognitive tasks.
- We allowed changes in the cutoffs if it increased the correlation at a 3-digit occupation level without a meaningful drop in the correlation at a 4-digit level, and if it mitigated any systematic differences between the task content measures calculated in PIAAC and STEP surveys.

Finally, in order to verify whether the values of task contents do not depend on the data source (PIAAC or STEP), we estimate a range of OLS regressions. In the base model, we regress (OLS) each task content measure against individual characteristics (gender, 10-year age groups, education, 1-digit occupations, sectors) and the STEP survey fixed effect which turns out negative and significant for all tasks except non-routine cognitive personal (Table B3). When we control for the level of literacy skills and GDP per capita,²⁶ the difference between STEP and PIAAC remains significant only in the case of manual tasks. This shows our survey measures of cognitive tasks are consistent and comparable between the two surveys. However, the STEP fixed effect remains significant even in the most elaborate specification. Therefore, we correct the values of manual task scores in STEP by this fixed effect (we add 0.17 to the manual task score of each individual in STEP sample).

	Non-routine cognitive	Non-routine	Routine	Manual
	analytical	cognitive personal	cognitive	Wallua
	Base model, total s	ample of 42 countries		
STEP dummy	-0.22***	-0.03	-0.05	-0.38***
Base mo	odel, subsample of 39 cou	ntries with literacy asses	sment data	
STEP dummy	-0.17**	-0.08	-0.17	-0.39***
Base model + controls for literacy	skills and for GDP per ca	pita, subsample of 39 cou	untries with literacy	assessment data
Literacy skills level: 0 and 1	-0.10***	-0.04***	-0.02	0.02
Literacy skills level: 3	0.08***	0.05***	-0.09***	-0.14***
Literacy skills level: 4 and 5	0.16***	0.11***	-0.22***	-0.30***
GDP per capita	-0.95	-1.51***	1.41	0.27
GDP per capita squared	0.05	0.08***	-0.07	-0.01
STEP dummy	-0.00	0.06	-0.07	-0.18***

Table B3. OLS regressions of task measures on sets of control variables and a STEP dummy

Note: the base regressions include dummies for gender, 10-year age groups, education, 1-digit occupations and sectors. To save space, we report only the coefficients for the STEP dummy, literacy skills and GDP per capita (in 1000s, in PPP, current international \$, country averages for 2011-2016). The regressions with literacy scores exclude China (CULS), Laos and Macedonia due to lack of literacy skills assessment in these countries. The total number of observations equals around 155,500 for the base model regression with all countries and around 144,500 for the specifications without China (CULS), Laos and Macedonia. The standard errors are clustered at a country level. Source: own estimations based on PIAAC, STEP, CULS and World Bank data.

²⁶ The literacy skills tests in STEP and PIAAC follow the same methodology and are comparable.

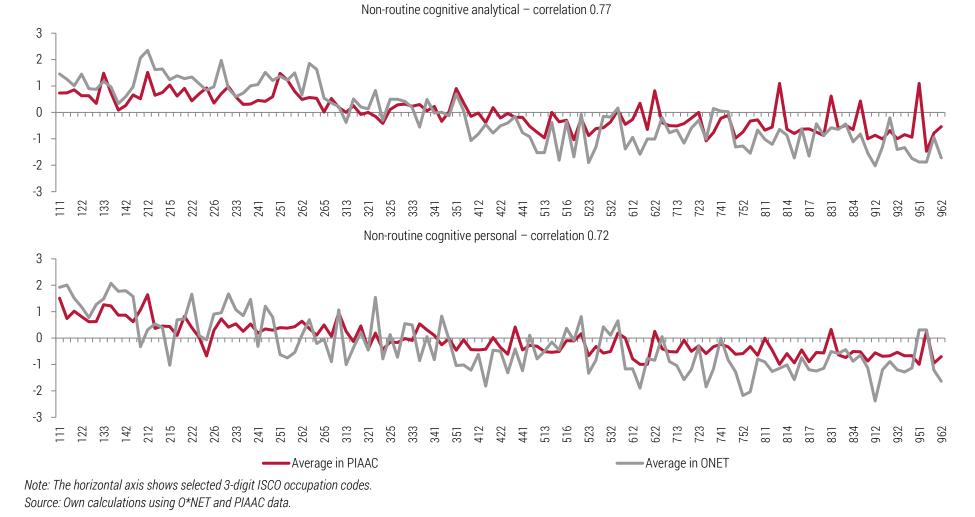


Figure B1. Values of task contents across 3-digit ISCO occupations in the United States.

39

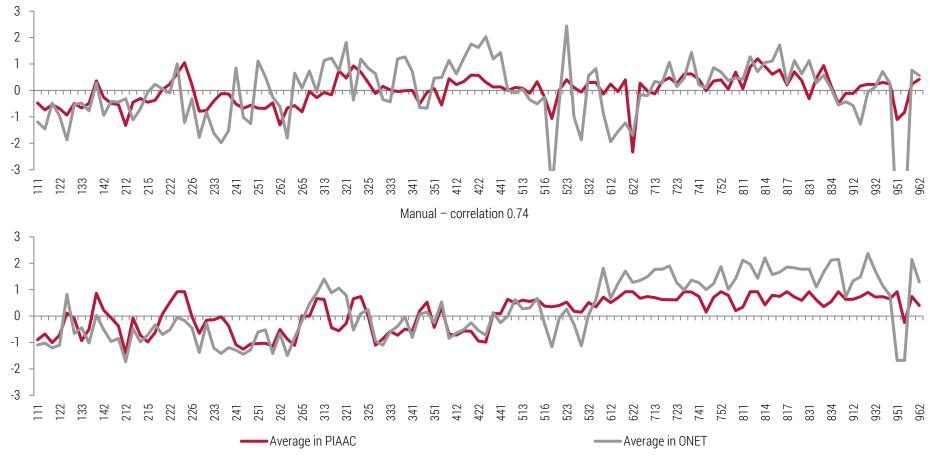


Figure B1. Values of task contents across 3-digit ISCO occupations in the United States (cont'd). Routine cognitive – correlation 0.55

Note: The horizontal axis shows selected 3-digit ISCO occupation codes. In order to use the same range for all tasks, the negative outliers in the O*NET routine cognitive tasks are truncated at -3: occupation 521 (Street and Market Salespersons) which has the value of -3.86, and occupation 951 (Street and Related Services Workers) and 952 (Street Vendors, excluding food) which have the value of -5.29.

Source: Own calculations using O*NET and US PIAAC data.

Appendix C. Relevant task items in PIAAC and STEP surveys

Task item name	e PIAAC	STEP				
	Question	Answers	Question Answe			
	In your job, how often do you	1. Never	As a regular part of this work, do you have to read			
	usually	2. Less than	the following?			
	 Read bills, invoices, bank 	once a		-		
Reading bills	statements or other	month	- Bills or financial statements			
	financial statements?	3. Less than				
	- Read articles in	once a		-		
Reading news	newspapers, magazines	week but at	- Newspapers or magazines			
	or newsletters?	least once a		Yes / No		
Reading	- Read articles in	month		-		
professional	professional journals or	4. At least	- Reports			
journals	scholarly publications?	once a				
Reading	- Read manuals or	week but		-		
manuals	reference materials?	not every	- Instruction manuals/operating manuals			
Filling forms	- Fill in forms?	day 5. Every day	As part of this work, do you fill out bills or forms?	-		
	In your job, how often do you	. ,	As a normal part of this work, do you do any of the			
	usually		following?			
	- Use more advanced math	•		-		
	or statistics such as					
Advanced	calculus, complex		- Use more advanced math, such as algebra,			
math	algebra, trigonometry or	A 1	geometry, trigonometry, etc.	Acchava		
	use of regression	As above		As above		
	techniques?					
Calculating	- Calculate prices, costs or		Calculata prices ar costa	-		
prices	budgets?		- Calculate prices or costs			
Calculating	- Use or calculate fractions,		- Use or calculate fractions, decimals or	•		
fractions	decimals or percentages?		percentages			
	In your job, how often do you					
Programming	usually use a programming	As above	Does your work as [OCCUPATION] require the use of	As above		
riogramming	language to program or write	AS above	software programming?	AS above		
	computer code?					
Making	How often does your job usually		As part of this work, do you have to make formal			
speeches or	involve making speeches or	As above	presentations to clients or colleagues to provide	As above		
giving	presentations in front of five or	10 0000	information or persuade them of your point of view?	//3 00000		
presentations	more people?		information of persuade them of your point of view:			
	And how often are you usually		Some tasks are pretty easy and can be done right	1. Never		
	confronted with more complex		away or after getting a little help from others. Other	2. Less th		
	problems that take at least 30		tasks require more thinking to figure out how they	once		
Solving problems	minutes to find a good solution?		should be done. As part of this work as	month		
	The 30 minutes only refers to	As above	[OCCUPATION], how often do you have to undertake	3. Less th		
problemo	the time needed to THINK of a		tasks that require at least 30 minutes of thinking	once		
	solution, not the time needed to		(examples: mechanic figuring out a car problem,	week but		
	carry it out.		budgeting for a business, teacher making a lesson	least once		
	ourry it out.		budgeting for a business, teacher making a lesson	month		

Table C1. The considered task items, their exact wordings and possible answers in PIAAC and STEP surveys

			plan, restaurant owner creating a new menu/dish for restaurant, dress maker designing a new dress)	 4. At least once a week but not every day 5. Every day
Physical tasks	How often does your job usually involve working physically for a long period?	As above	Using any number from 1 to 10 where 1 is not at all physically demanding (such as sitting at a desk answering a telephone) and 10 is extremely physically demanding (such as carrying heavy loads, construction worker, etc.), what number would you use to rate how physically demanding your work is?	1-10
Supervising	Do you manage or supervise other employees?	Yes / No	As a normal part of this work do you direct and check the work of other workers (supervise)?	Yes / No
Collaborating	In your job what proportion of your time do you usually spend cooperating or collaborating with co-workers?	 None of the time Up to a quarter of the time Up to half of the time More than half of the time All the time 	As part of this work, how frequently do you spend time co-operating or collaborating with co-workers?	 Never Less than once a month Less than once a week but at least once a month At least once a week but not every day Every day
Changing order of tasks	The next few questions are about the amount of flexibility you have in deciding how you do your job: To what extent can you choose or change the sequence of your tasks?	 Not at all Very little To some extent To a high extent To a very high extent 	Still thinking of your work as [OCCUPATION] how much freedom do you have to decide how to do your work in your own way, rather than following a fixed procedure or a supervisor's instructions? Use any number from 1 to 10 where 1 is no freedom and 10 is complete freedom.	1-10

Note: the PIAAC questions wordings in this table come from the International Master Questionnaire, available at the OECD website.²⁷ The STEP questions wordings in this table come from the English version of the Armenia STEP Skills Measurement Survey, available at the World Bank's microdata website.²⁸

Source: own elaboration based on PIAAC and STEP.

²⁷ See www.oecd.org/skills/piaac/BQ_MASTER.HTM [accessed: 2017-05-02].

²⁸ See microdata.worldbank.org/index.php/catalog/2010 [accessed: 2017-05-04].

To ensure comparability between STEP and PIAAC data, we rescale the answers to achieve common answer scales in both surveys. The PIAAC questions typically refer to the frequency of performing a task (five levels ranging from 'never' to 'every day'), while many STEP questions refer to whether the responders normally perform a specific task as part of their job or not. Out of 16 questions we consider, two have five available answers in both PIAAC and STEP, and two have 'Yes/No' answers in both PIAAC and STEP. For these questions, we use original variables. For 10 questions which have five available answers in PIAAC but a 'Yes/No' answer in STEP, we convert PIAAC variables into four variants of dummy variables based on the cuttoffs in original answers. For two questions which have five available answers in PIAAC and 10 available answers in STEP, we recode the STEP answers into a 1 to 5 scale (1 and 2 to 1, 3 and 4 to 2,..., 9 and 10 to 5). We also correct the item indicating supervising other workers in the STEP data so that only individuals with co-workers are allowed to supervise others.²⁹ In the PIAAC data all of the self-employed responders who had no other workers in their jobs indicated they did not supervise anyone. Since this item has a consistent wording in both surveys, our correction of values in STEP ensures consistency with PIAAC data.

²⁹ Some respondents in STEP indicated supervising other workers despite declaring that they worked alone. Our change corrects this in cases where respondents indicated any of the following combinations: a) being self-employed with no hired workers, b) being self-employed with no unpaid or paid workers, c) being the only paid worker at the current job or that the total number of people working at the organization equals one (the respondent). This problem is not present in CULS.

Appendix D. Comparison of task measures based on STEP and PIAAC data

Table D1. Comparison of task measures based on STEP and PIAA	C data
--	--------

Our measures; PIAAC and STEP		Lo Be	ello et al. (2019); STEP	de la	Rica et al. (2020); PIAAC	Marcolin et al. (2019); PIAAC; Routine Intensity Index only	
Task content	Items	Task content	Items	Task content Items		ltems	
Reading news Non-routine Reading professional journals cognitive analytical Solving problems Programming		Non-routine analytical	Number of types of documents read Length of longest documents typically read Solving problems Number of math tasks performed		Advanced math Planning the activities of others Solving problems Persuading or influencing people		
Non-routine cognitive personal	Supervising Making speeches or giving presentations	Non-routine interpersonal	Supervising Contact with clients	– Rea	Read professional journals		
Routine cognitive	Changing order of tasks (reversed) Filling forms Making speeches or giving presentations (reversed)	Routine cognitive	Learning new things Autonomy Repetitiveness	Routine	Planning your own activities (reversed) Organising own time (reversed) Instructing, training or teaching (reversed) Advising people (reversed) Making speeches or giving presentations (reversed)	Planning own activities Changing order of tasks Change how to do work Organising own time	
		Non-routine manual	Driving Repair electronic equipment			-	
Manual	Physical tasks	Routine manual	Operate machines or equipment Physical tasks		Using skill or accuracy with hands or fingers Physical tasks	-	
Methods: uniform coding in STEP and PIAAC; standardisation (means and standard deviations); averages		Methods: standardization (means and standard deviations); summation		Methods: principal component analysis		Methods: averages.	

Source: own elaboration based on Lo Bello et al. (2019), de la Rica et al. (2020) and Marcolin et al. (2019).

Appendix E. Other data sources

In order to estimate the cross-country regressions, we merge the PIAAC, STEP and CULS data with three additional variables: ICT stock per worker, number of robots per worker, and the global value chain participation.

The data on ICT capital stock come from Eden and Gaggl (2020). The data are available at the country level, except eight countries in our sample: Armenia, Cyprus, Georgia, Ghana, Estonia, Kazakhstan, Laos, and Macedonia. The latest year available is 2011.

The data on robots come from the International Federation of Robotics [IFR] (2017). The latest data available are from 2016 but we use the average for 2011-2016 since our survey data cover this period. The IFR data are available for ISIC 4 sectors: A, B, C, D and E (jointly), F and P. We aggregate them to three broad categories: Agriculture, Industry and Services and calculate the number of robots per worker in each country / sector cell. The IFR data are unavailable for 12 countries in our sample: Armenia, Bolivia, Cyprus, Ecuador, Georgia, Ghana, Kazakhstan, Kenya, Laos, Macedonia, Mexico, and Peru.

The data on global value chain participation are sourced from the UIBE GVC INDICATORS database. We use the backward linkage-based measure, defined as the foreign value added share in production of final goods and services, and the forward-linkage measure, defined as the domestic value added from production of intermediate exports or domestic factor content in intermediate exports (Wang et al., 2017). We use the variables based on GTAP. The latest year available is 2011. We merge the UIBE GVC INDICATORS data with our data at the country-industry level. As the sector classifications are not fully compatible, we aggregate some of the ISIC 4 categories to broader groups: "E+O+P+Q+U" (water supply; sewerage, waste management and remediation activities; public administration and defense; compulsory social security; education; human health and social work activities; activities of extraterritorial organizations and bodies). In China (CULS) this group also includes category D (electricity, gas, steam and air conditioning supply). "G+I" (wholesale and retail trade, repair of motor vehicles and motorcycles; accommodation and food service activities); "L+M+N" (real estate activities; professional, scientific and technical activities; activities of households as employers; undifferentiated goods- and services-producing activities of households for own use). The UIBE GVC INDICATORS data are not available for Macedonia.

Appendix F. Additional regression results

Table F1. The estimated interaction terms between sector fixed effects and GDP per capita (log, demeaned), benchmark specification as in Table 4

	All workers	High-skilled occupations (ISCO 1-3)	Middle-skilled occupations (ISCO 4-5)	Low-skilled occupations (ISCO 7-9)
A	-0.132**	-0.047	-0.010	-0.199***
Agriculture [A]	(0.055)	(0.070)	(0.104)	(0.063)
Mining [D]	-0.127**	-0.057	-0.038	-0.241***
Mining [B]	(0.055)	(0.074)	(0.084)	(0.075)
Manufacturing [0]	0.014	0.001	0.054	-0.023
Manufacturing [C]	(0.051)	(0.057)	(0.059)	(0.053)
Electricity & Water supply	-0.116*	-0.136*	-0.093	-0.066
[D+E]	(0.062)	(0.071)	(0.132)	(0.058)
Construction [[]	-0.083*	-0.037	-0.062	-0.164***
Construction [F]	(0.049)	(0.050)	(0.082)	(0.053)
Transportation and storage	0.010	0.015	-0.018	-0.044
[H]	(0.050)	(0.057)	(0.071)	(0.053)
Accommodation and food	-0.061	0.001	-0.052	-0.063
service [I]	(0.052)	(0.069)	(0.061)	(0.058)
Information and	0.051	0.064	0.084	0.029
communication [J]	(0.057)	(0.064)	(0.063)	(0.080)
	-0.018	-0.034	0.043	-0.060
Financial and insurance [K]	(0.064)	(0.056)	(0.069)	(0.112)
Real estate & Professional	-0.008	0.147*	-0.103	-0.026
[L]	(0.064)	(0.084)	(0.069)	(0.184)
Administrative [MJN]	0.055	0.037	0.061	0.080
Administrative [M+N]	(0.051)	(0.062)	(0.048)	(0.056)
Dublic administration [0]	0.062	0.073	0.018	0.058
Public administration [0]	(0.062)	(0.057)	(0.080)	(0.101)
Education [D]	0.098*	0.082	0.004	0.010
Education [P]	(0.050)	(0.052)	(0.083)	(0.059)
Llumon boolth [0]	0.010	-0.009	-0.057	-0.021
Human health [Q]	(0.047)	(0.049)	(0.062)	(0.078)
Arto [D]	-0.019	-0.041	0.029	-0.157
Arts [R]	(0.063)	(0.057)	(0.060)	(0.098)
Other convice [S]	-0.062	-0.142**	0.009	-0.028
Other service [S]	(0.052)	(0.058)	(0.052)	(0.056)
Activities of household [T]	-0.010	0.368	0.060	-0.099
	(0.078)	(0.371)	(0.059)	(0.088)
Extraterritorial organizations	0.150*	0.127	-0.126	0.313
[U]	(0.083)	(0.089)	(0.180)	(0.230)
No. of observations	166,542	68,027	52,906	45,609
R-squared	0.222	0.115	0.088	0.084

*Note: *** p<0.01, ** p<0.05, * p<0.1. Standard errors in parentheses. We use standardized weights that give each country equal weight. Source: own estimations based on PIAAC, STEP, CULS World Bank, and UIBE GVC Indicators data.*

Table F2. The correlates of routine task intensity (RTI) at the worker level, including forward linkage-based measure of participation in global value chains (domestic value added from production of intermediate exports or domestic factor content in intermediate exports), OLS

	All workers	High-skilled occupations (ISCO 1-3)	Middle-skilled occupations (ISCO 4-5)	Low-skilled occupations (ISCO 7-9)
Computer use	1.642***	0.843**	0.846*	2.004***
	(0.346)	(0.357)	(0.474)	(0.423)
Computer use ^ 2	-2.164***	-1.410***	-1.405***	-2.437***
	(0.291)	(0.307)	(0.414)	(0.374)
Global Value Chain (GVC)	0.296***	0.073	0.505***	0.477***
Participation (forward linkage-based)	(0.112)	(0.113)	(0.175)	(0.158)
Ln(GDP per capita) –	0.041	-0.012	0.040	0.096*
Mean(Ln(GDP per capita)	(0.039)	(0.043)	(0.052)	(0.050)
GVC participation *	-0.261***	-0.157	-0.148	-0.093
[Ln(GDP pc) –mean(Ln(GDP pc)]	(0.097)	(0.110)	(0.144)	(0.139)
FDI / GDP	0.005	0.019***	0.001	-0.010
	(0.005)	(0.006)	(0.007)	(0.007)
Education: primary	0.284***	0.150***	0.263***	0.158***
	(0.015)	(0.026)	(0.018)	(0.020)
Education: tertiary	-0.500***	-0.274***	-0.220***	-0.179***
	(0.016)	(0.016)	(0.019)	(0.034)
Literacy skills level:	0.100***	0.033	0.058**	0.088***
1 or lower	(0.016)	(0.023)	(0.024)	(0.021)
Literacy skills level: 3	-0.131***	-0.094***	-0.046**	-0.045**
	(0.012)	(0.014)	(0.019)	(0.022)
Literacy skills level:	-0.270***	-0.192***	-0.038	-0.164***
4 and 5	(0.018)	(0.018)	(0.027)	(0.041)
Female	0.239***	0.223***	0.197***	0.338***
	(0.011)	(0.012)	(0.017)	(0.023)
Age: 16-24	0.202***	0.213***	0.182***	0.122***
	(0.016)	(0.025)	(0.023)	(0.020)
Age: 35-44	-0.057***	-0.057***	-0.028*	-0.056***
	(0.010)	(0.013)	(0.016)	(0.019)
Age: 45-54	-0.023*	-0.058***	0.006	0.010
	(0.012)	(0.014)	(0.020)	(0.019)
Age: 55-65	0.020	-0.042**	0.096***	0.059***
	(0.015)	(0.018)	(0.023)	(0.022)
Sector fixed effects	Yes	Yes	Yes	Yes
No. of observations	166,542	68,027	52,906	45,609
R-squared	0.223	0.115	0.088	0.081

Note: *** p<0.01, ** p<0.05, * p<0.1. Standard errors in parentheses. We use standardized weights that give each country equal weight. The reference levels are: age 25-34, secondary education, wholesale and retail trade; repair of motor vehicles and motorcycles (ISIC G), lower medium literacy skills (level 2).

regressions	Model		d ICT capital as puntries)	controls	Benchmark specification for 32 countries with robot and ICT data available			
	All workers	High-skilled occupation s (ISCO 1-3)	Middle- skilled occupations (ISCO 4-5)	Low-skilled occupation s (ISCO 7-9)	All workers	High-skilled occupation s (ISCO 1-3)	Middle- skilled occupation s (ISCO 4-5)	Low-skilled occupation s (ISCO 7-9)
Computer use	1.342**	0.967*	0.713	1.151**	1.593***	1.354**	0.860	1.328***
·	(0.528)	(0.549)	(0.697)	(0.505)	(0.514)	(0.531)	(0.658)	(0.501)
Computer use ^2	-1.841***	-1.250***	-1.409**	-1.757***	-2.065***	-1.565***	-1.554***	-1.959***
	(0.414)	(0.443)	(0.557)	(0.437)	(0.405)	(0.430)	(0.522)	(0.437)
ICT stock per	-0.045***	-0.058***	-0.032*	-0.044**				
worker	(0.013)	(0.018)	(0.019)	(0.022)				
Robots per	-0.001	-0.019	-0.042***	-0.009				
worker	(0.010)	(0.012)	(0.016)	(0.010)				
FVA share	0.194	-0.220	0.282	0.687***	0.224	-0.180	0.305	0.719***
	(0.171)	(0.169)	(0.188)	(0.200)	(0.174)	(0.178)	(0.195)	(0.199)
Ln(GDP per capita) –	0.112	0.005	0.132	0.194	0.066	-0.064	0.104	0.148*
mean(Ln(GDP per capita))	(0.074)	(0.061)	(0.092)	(0.171)	(0.070)	(0.053)	(0.085)	(0.078)
FVA share * [Ln(GDP pc) –	-0.266	-0.386*	-0.218	-0.077	-0.283	-0.397*	-0.239	-0.093
mean(Ln(GDP pc)]	(0.199)	(0.227)	(0.222)	(0.227)	(0.200)	(0.232)	(0.226)	(0.227)
	0.028*	0.071***	0.049**	-0.010	0.038**	0.086***	0.058***	0.001
FDI / GDP	(0.015)	(0.016)	(0.023)	(0.023)	(0.015)	(0.016)	(0.022)	(0.023)
Education:	0.306***	0.152***	0.249***	0.164***	0.305***	0.151***	0.249***	0.163***
primary	(0.016)	(0.029)	(0.020)	(0.019)	(0.016)	(0.029)	(0.020)	(0.019)
Education:	-0.518***	-0.293***	-0.218***	-0.200***	-0.517***	-0.294***	-0.218***	-0.197***
tertiary	(0.016)	(0.017)	(0.020)	(0.036)	(0.016)	(0.017)	(0.020)	(0.037)
Literacy skills level:	0.161***	0.088***	0.095***	0.130***	0.161***	0.088***	0.094***	0.130***
1 or lower	(0.017)	(0.032)	(0.028)	(0.023)	(0.017)	(0.032)	(0.028)	(0.023)
Literacy skills	-0.133***	-0.091***	-0.043**	-0.055**	-0.134***	-0.092***	-0.044**	-0.055**
level: 3	(0.013)	(0.016)	(0.021)	(0.023)	(0.013)	(0.016)	(0.021)	(0.023)
Literacy skills	-0.254***	-0.179***	-0.034	-0.161***	-0.254***	-0.180***	-0.033	-0.160***
level: 4 and 5	(0.018)	(0.021)	(0.029)	(0.040)	(0.018)	(0.021)	(0.029)	(0.040)
Female	0.259***	0.226***	0.230***	0.378***	0.260***	0.227***	0.233***	0.379***
	(0.013)	(0.012)	(0.019)	(0.029)	(0.013)	(0.012)	(0.020)	(0.029)
Ado: 16.04	0.267***	0.264***	0.255***	0.145***	0.265***	0.262***	0.254***	0.145***
Age: 16-24	(0.017)	(0.028)	(0.027)	(0.022)	(0.017)	(0.028)	(0.027)	(0.022)
Ago: 25 44	-0.095***	-0.077***	-0.069***	-0.087***	-0.094***	-0.077***	-0.068***	-0.088***
Age: 35-44	(0.011)	(0.013)	(0.020)	(0.019)	(0.011)	(0.013)	(0.020)	(0.019)
Age: 45-54	-0.084***	-0.087***	-0.047**	-0.040**	-0.083***	-0.087***	-0.044**	-0.039**
Ayt. 4J-04	(0.012)	(0.016)	(0.022)	(0.020)	(0.012)	(0.016)	(0.022)	(0.020)

Table F3. The correlates of RTI, including ICT stock and robot stock per worker, 32 countries with available data, OLS regressions

-0.052***	-0.083***	0.058**	-0.004	-0.052***	-0.084***	0.060**	-0.003
(0.015)	(0.018)	(0.028)	(0.023)	(0.015)	(0.018)	(0.028)	(0.023)
Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
124,046	54,927	36,422	32,697	124,046	54,927	36,422	32,697
0.240	0.119	0.100	0.108	0.239	0.117	0.099	0.107
	(0.015) Yes 124,046	(0.015) (0.018) Yes Yes 124,046 54,927	(0.015) (0.018) (0.028) Yes Yes Yes 124,046 54,927 36,422	(0.015) (0.018) (0.028) (0.023) Yes Yes Yes Yes Yes 124,046 54,927 36,422 32,697	(0.015) (0.018) (0.028) (0.023) (0.015) Yes Yes Yes Yes Yes Yes 124,046 54,927 36,422 32,697 124,046	(0.015) (0.018) (0.028) (0.023) (0.015) (0.018) Yes Yes Yes Yes Yes Yes Yes Yes 124,046 54,927 36,422 32,697 124,046 54,927	(0.015) (0.018) (0.028) (0.023) (0.015) (0.018) (0.028) Yes Ye

Note: *** p<0.01, ** p<0.05, * p<0.1. Standard errors in parentheses. We use standardized weights that give each country equal weight. The reference levels are: age 25-34, secondary education, middle-skilled occupations (ISCO 4-5), wholesale and retail trade; repair of motor vehicles and motorcycles (ISIC G), lower medium literacy skills (level 2). The coefficients for sector fixed effects are not presented in order to save space, and are available on request. ICT stock per worker, robots per worker, FDI and the FVA share in domestic production variables are standardized in our sample.

Source: own estimations based on PIAAC, STEP, CULS, World Bank, Eden and Gaggl (2020) and IFR and UIBE GVC Indicators data.

		All workers			
Computer use	-0.571*	Formala	0.198	Sector J	-0.053
	(0.341)	Female	(0.176)		(0.047)
Computer use ^2	-0.741**	A 16 04	-0.207	Sector J	-0.228***
	(0.312)	Age: 16-24	(0.206)		(0.077)
	0.246**		0.189	Sector K	-0.071
Foreign Value Added share	(0.123)	Age: 35-44	(0.166)		(0.060)
Ln(GDP per capita) –	0.088**		0.277	Sector L	-0.086*
mean(Ln(GDP per capita))	(0.035)	Age: 45-54	(0.202)		(0.050)
FVA share *	-0.121		-0.584**	Sector M+N	-0.116***
[Ln(GDP pc) –mean(Ln(GDP pc)]	(0.100)	Age: 55-65	0.198		(0.037)
FDI / GDP	-0.008	Sector A	-0.049	Sector O	-0.110**
	(0.006)		(0.056)		(0.049)
Education: primary	0.041	Sector B	-0.250***	Sector P	-0.493***
	(0.161)		(0.079)		(0.061)
Education: tertiary	0.120	Sector C	-0.094**	Sector Q	-0.166**
	(0.138)		(0.042)		(0.065)
Literacy skills level:	-0.174	Sector D+E	-0.079	Sector R	-0.285***
1 or lower	(0.200)		(0.073)		(0.046)
Literacy skills level: 3	-0.016	Sector F	-0.255***	Sector S	-0.290***
	(0.213)		(0.078)		(0.051)
Literacy skills level:	-0.700**	Sector H	0.037	Sector T	0.138
4 and 5	(0.294)		(0.064)		(0.121)
		Sector I		Sector U	-0.113
					(0.116)
No. of observations			820		
R-squared			0.806		

Table F4. The correlates of routine task intensity (RTI) at the sector level, OLS

Note: *** p<0.01, ** p<0.05, * p<0.1. Standard errors in parentheses. All variables are calculated averages in sector *s* in country *c*. We use standardised weights that give each country equal weight. The reference levels are: age 25-34, secondary education, wholesale and retail trade; repair of motor vehicles and motorcycles (ISIC G), lower medium literacy skills (level 2). RTI, FDI and the FVA share in domestic production variables are standardized in our sample.

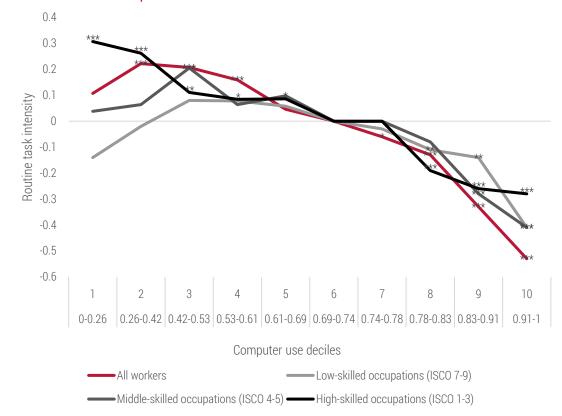
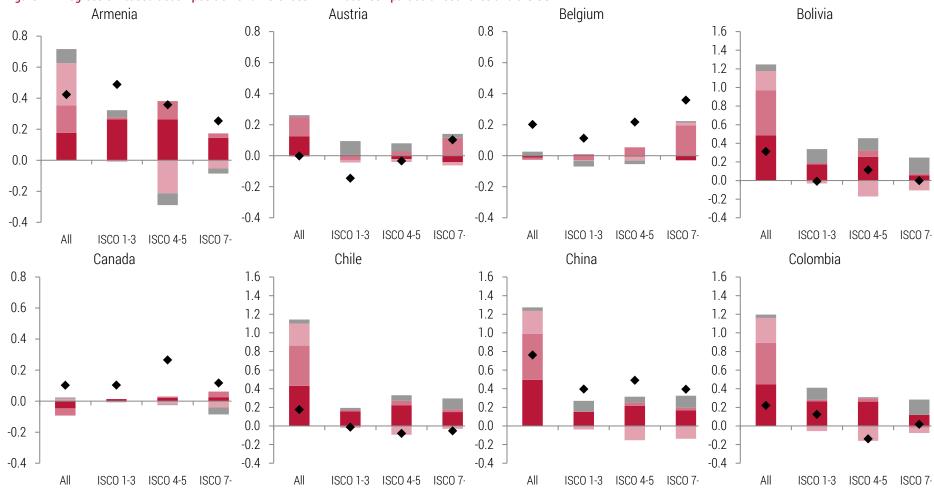


Figure F1. Estimated relationship between computer use and RTI, for all workers and by occupational group, using the deciles of sector shares of computer use.

Notes: *** p<0.01, ** p<0.05, * p<0.1. Estimates from the baseline specification (2) (presented in Table 4) but with decile computer use fixed effects used instead of a continous computer use variable. The 6th decile of a pooled distributon of a country-sector level computer use shares forms the reference group in all regressions. Full estimation results are available upon request. Source: own estimations based on PIAAC, STEP, CULS, World Bank and UIBE GVC Indicators data.



Appendix H. Decomposition results for all countries

Figure H1. Regression-based decomposition of differences in RTI between particular countries and the US.

Note: Results of decomposition (3) based on regressions presented in Table 4. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US. Source: own estimations based on PIAAC, STEP, CULS, World Bank and UIBE GVC Indicators data.

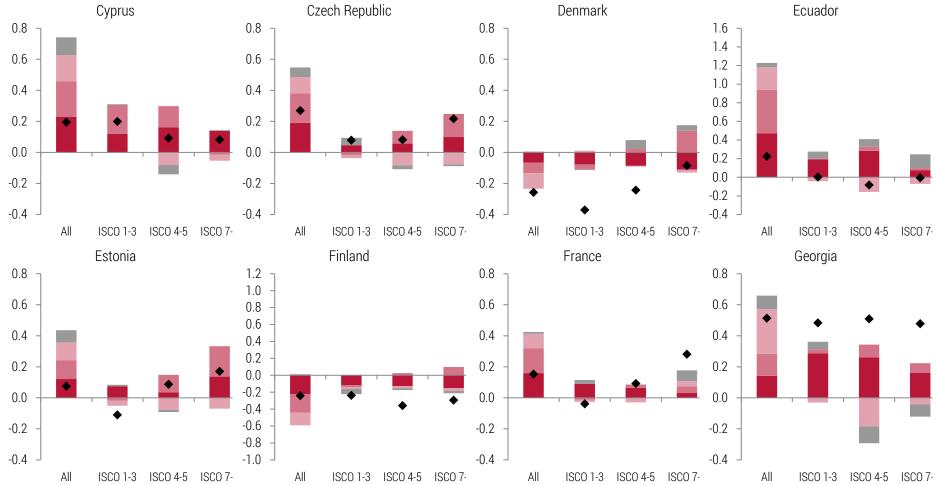


Figure G1 (cont'd). Regression-based decomposition of differences in RTI between particular countries and the US.

Note: Results of decomposition (3) based on regressions presented in Table 4. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US. Source: own estimations based on PIAAC, STEP, CULS, World Bank and UIBE GVC Indicators data.

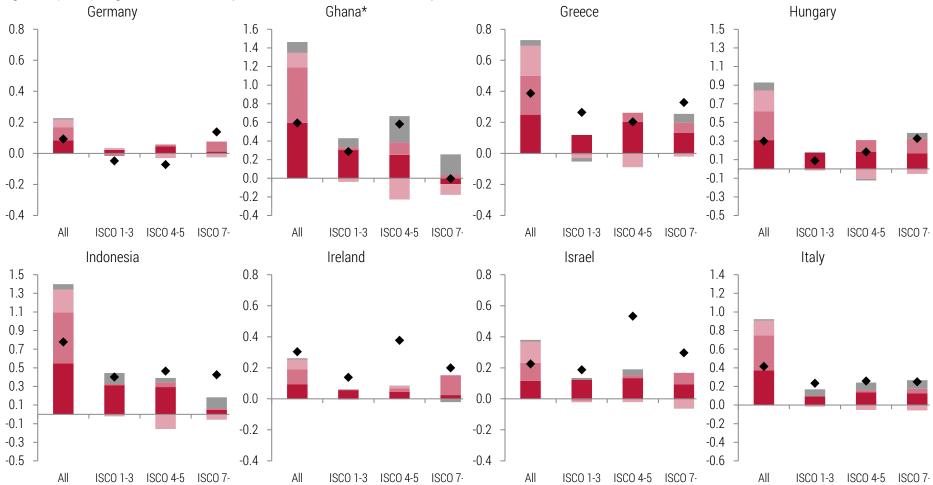


Figure G1 (cont'd). Regression-based decomposition of differences in RTI between particular countries and the US.

Note: Results of decomposition (3) based on regressions presented in Table 4. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US. * Figures for countries marked with asterisk have different vertical scale than most countries.

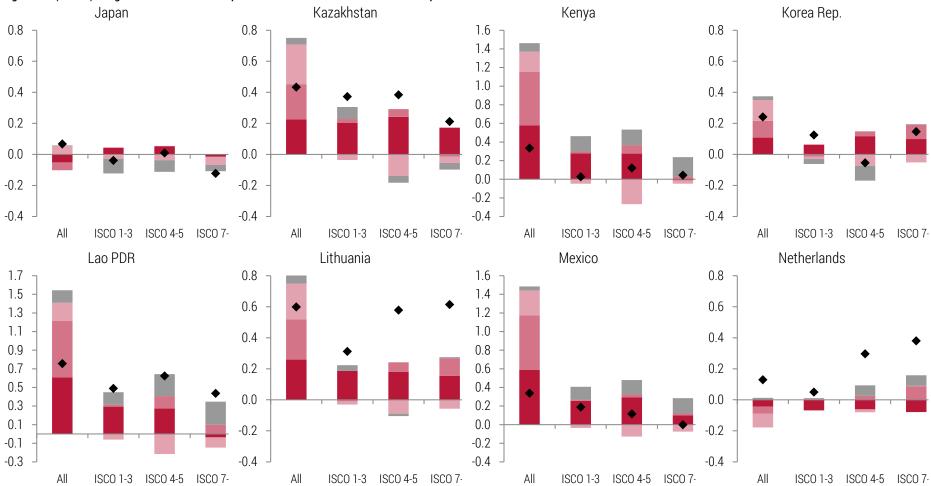


Figure G1 (cont'd). Regression-based decomposition of differences in RTI between particular countries and the US.

Note: Results of decomposition (3) based on regressions presented in Table 4. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US. * Figures for countries marked with asterisk have different vertical scale than most countries.



Figure G1 (cont'd). Regression-based decomposition of differences in RTI between particular countries and the US.

Note: Results of decomposition (3) based on regressions presented in Table 4. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US. * Figures for countries marked with asterisk have different vertical scale than most countries.

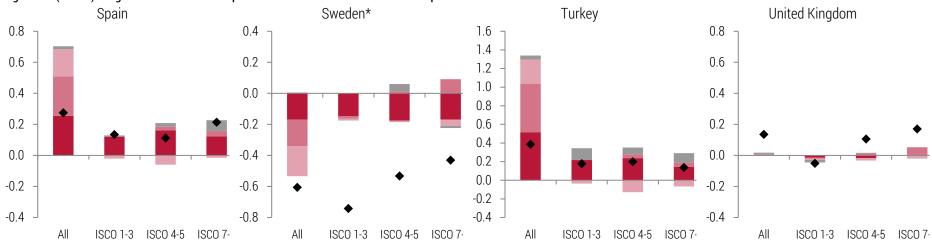


Figure G1 (cont'd). Regression-based decomposition of differences in RTI between particular countries and the US.

Note: Results of decomposition (3) based on regressions presented in Table 4. 0 is set at the US average value and 1 corresponds to one standard deviation of RTI in the US. * Figures for countries marked with asterisk have different vertical scale than most countries.



www.ibs.org.pl